Diamond Cuts Diamond: News Co-mention Momentum Spillover Prevails in China

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Abstract

We conduct a comprehensive study on momentum spillovers in the Chinese stock market using various types of economic linkages. By developing a exible and innovative algorithm to identify linkages among listed rms using millions of Chinese business news articles, we nd that the news co-mention momentum spillover is signi cantly stronger compared to other forms of momentum spillovers. Using spanning tests and Fama-MacBeth regressions, we further show that the news co-mention momentum spillover uni es all di erent forms of momentum spillover e ects in the Chinese stock market. Notably, the analyst co-coverage momentum spillover e ect in the Chinese stock market. We further explore the di erences in the information content of links implied by news co-mentioning and other proxies. We suggest that the dominance of news co-mention momentum spillover over others can be attributed to two primary factors: comprehensive information and prompt updates.

Keywords: Economic Linkage, Big Textual Data, Momentum Spillover, News co-mention, Limited attention

JEL Classi cation: G11, G12, G14

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1 Introduction

In the past decade, a new momentum-like anomaly known as \momentum spillovers" or \cross- rm momentum" has raised a lot of attention in the US stock market. A focal rm's stock price responds sluggishly to new information about economically linked rms, and as a result, the past returns of linked rms predict the future return of the focal rm. The literature documents the existence of such momentum spillover e ect via various types of economic linkages, including industry links (Moskowitz and Grinblatt, 1999; Hou, 2007), customer-supply links (Cohen and Frazzini, 2008), geographic links (Parsons et al., 2020; Jin and Li, 2020), technology links (Lee et al., 2019; Duan et al., 2022), news-implied links (Scherbina and Schlusche, 2013), concept links (Du et al. (2022)), and analyst co-coverage links (Ali and Hirshleifer, 2020), etc.

In the US market, the cross- rm momentum via analyst co-coverage uni es the \zoo" of other momentum spillover e ects (Ali and Hirshleifer, 2020). Unlike developed markets, the Chinese stock market is relatively young, and it possesses several distinctive features that make it particularly intriguing for research. The rst prominent feature of the Chinese stock market is the dominance of retail investors who are less sophisticated (Bailey et al., 2009). It's generally more costly for them to collect and process information, and they are more likely to be inattentive to new information about linked rms while making investing decisions. Consequently, we might expect a stronger lead-lag e ect among linked rms in the Chinese stock market. The second prominent feature of the A market is that investors trade in response to news quickly (Pan et al., 2016), which leads to its high turnover rate. Given these special features, two natural questions arise: whether various forms of momentum spillover e ects exist in the Chinese stock market and whether there exists a dominant species that uni es the \zoo" of all momentum spillover e ects.

In this paper, we conduct a comprehensive study on momentum spillovers in the Chinese stock market using various types of economic linkages. We nd that the momentum spillover e ect among rms linked through news co-mentioning (referred to as the news co-mention momentum spillover) is signi cantly stronger than others. Using spanning tests and Fama and MacBeth (1973) regressions, we show that all di erent forms of cross- rm momentum e ects can be uni ed by the news co-mention momentum spillover in the Chinese stock market. Notably, the analyst co-coverage momentum spillover e ect, which is the dominant species in the US stock market, is subsumed by the news co-mention momentum spillover e ect in the Chinese stock market. We further explore the di erences in the information content of links implied by news co-mentioning and other proxies. And we suggest that the dominance of news co-mention momentum spillover over others can be attributed to two primary factors: comprehensive information and prompt updates.

To be speci c, to construct the news co-mention linkages, we adopt two strategies. The rst one follows Scherbina and Schlusche (2013) and Ge et al. (2022), which de ne two rms as linked if they are co-mentioned in the same piece of news article during a pre-speci ed identi cation window. This identi cation strategy is referred to as *same_article*. One potential issue with this identi cation strategy is that two rms that happen to appear in the same article may appear in di erent sentences and are actually unrelated to each other. Given this concern, the second identi cation strategy we adopt is called *same_sentence*, which de nes two rms as linked if they are co-mentioned in the same sentence of the piece of news article during a pre-speci ed identi cation window.¹

With the news-implied linkages, we proceed to test the existence of the news co-mention momentum spillover

¹A similar link identi cation strategy is adopted in Schwenkler and Zheng (2019) and Schwenkler and Zheng (2021).

e ect. Duan et al. (2022) document that the momentum spillover e ects in the A market are more prominent for the weekly returns as the turnover rate is extremely high in the Chinese stock market. In addition, He et al. (2021) argue that Chinese investors may pay more attention to the past week's returns rather than the past month's returns. Given the above reasons, we mainly focus on momentum spillovers at the weekly frequency. At the end of each week, we construct a signal for each stock based on the weighted average returns of its linked stocks. We sort stocks into quintiles, and a long-short strategy could be constructed, which involves the purchase of stocks from the top quintile while shorting stocks from the bottom quintile. We construct eight such long-short trading strategies using combinations of the two identic cation strategies and four identic cation windows (window lengths including 3-, 6-, 9-, and 12-month),² and these trading strategies that are based on news co-mention momentum spillover all generate signi cant and positive abnormal returns. For instance, adopting same_sentence identi cation strategy and a 3-month identi cation window, the long-short return is 1.94% (t-statistic=5.33), and the corresponding Liu et al. (2019) four-factor (CH-4 for short later) adjusted alpha is 1.86% (t-statistic=5.23).³ Moreover, we nd that the *same_sentence* strategy works better than the same_article strategy as the former yields larger abnormal returns with the same identi cation window. This nding aligns with our expectation that the *same_sentence* strategy is more e ective in identifying genuine links among rms. Another interesting nding is that the strength of news co-mention momentum spillover decreases steadily as the length of the identi cation window increases. This can be attributed to the fast updating of news, rendering stale information less useful in identifying relevant linkages.

We then compare the strength of news co-mention momentum spillover with various other momentum spillover e ects studied in the existing literature. The seven other cross- rm momentum e ects we examine include shared-analyst momentum, industry momentum, geographic momentum at the province and city levels, customer-supplier momentum, technology momentum, statistical momentum, and concept momentum. Similar to the previous case, for each linkage type, we construct signals for each stock based on the weighted average returns of its linked stocks and sort all stocks into quintiles. The long-short strategies exploiting most of those cross- rm momentum e ects also yield positive and statistically signi cant abnormal returns. However, these returns are noticeably smaller in both economic and statistical magnitudes compared to those obtained using news co-mentioning.

Given that the news co-mention momentum spillover e ect is signi cantly stronger than others, we follow Ali and Hirshleifer (2020) to use spanning tests and Fama and MacBeth (1973) regressions to investigate whether the news co-mention momentum spillover encompasses and uni es the various other momentum spillover e ects observed in the Chinese stock market.

To conduct the spanning tests, we construct the long-short factor returns based on news co-mentioning and the other six types of linkages discussed earlier.⁴ Considering the superior performance of the *same_sentence* identi cation strategy, we focus on this strategy for the construction of the news co-mention momentum spillover factor. We construct four factors using identi cation windows of 3, 6, 9, and 12 months. For instance, we denote Sentence_3 as the news co-mention momentum spillover factor based on identifying linkages using a 3-month window with the *same_sentence* strategy. Strikingly, the alphas of all other momentum spillover factors become

²The identi cation window cannot be too short as it would result in only a small number of rms with news-links.

³These returns are monthly returns converted from weekly ones.

⁴Due to the low quality of disclosed data regarding top customers in China, we do not consider the customer-supplier momentum spillover factor, which is often used as the representative cross- rm momentum in the US (e.g., Huang et al. (2021) and Huang et al. (2022)).

insigni cant, and some even turn negative when we add the news co-mention momentum spillover factor to the CH-4 factor model. For instance, the long-short strategy based on statistical momentum generates a CH-4 alpha of 1.34% (t-statistic=3.35), and its economic magnitude ranks second among all competitors. However, when we add the Sentence_3 factor, its CH-4+Sentence_3 model adjusted alpha decreases drastically to 0.17% (t-statistic=0.43). In contrast, none of the other six momentum factors can explain any of the four news co-mention momentum spillover factors (i.e., Sentence_3, Sentence_6, Sentence_9, Sentence_12). Additionally, among the four news co-mention momentum spillover factors, the factor constructed with a shorter identi cation window can explain the factors with longer identi cation windows, but not vice versa. The factor-spanning tests show that all di erent forms of cross- rm momentum e ects can be uni ed by the news co-mention momentum spillover in the Chinese stock market. The conclusion remains to be robust when we change the factor model, industry classi cation, and news co-mention type. Fama-MacBeth regressions further support this conclusion. Based on both union and intersection samples, after adding the news co-mention peer rm returns as an explanatory variable, the variables based on other cross- rm momentum e ects all become insigni cant.

We then explore the di erences in the information content of links implied by news co-mentioning and other proxies. Scherbina and Schlusche (2013), Schwenkler and Zheng (2019) nd that news co-mentioning reveals a wide range of economic linkages among companies, including same industry, business alliances, partnerships, banking and nancing, customer-supplier, production similarity, etc. To examine whether the news co-mention momentum spillover is stronger than others because news-implied linkages are more comprehensive, we conduct an exercise to analyze the within- and cross-industry momentum spillovers separately for each type of linkage. Firstly, we nd that, for the news co-mention linkage, the degree of overlap with industry information is relatively low compared to other types of linkages. On the other hand, the shared-analyst linkage shows the highest degree of overlap with the industry linkage. Next, for each linkage type, we decompose linked rms into industry peers and cross-industry peers, and we examine the two corresponding momentum spillover e ects separately. We nd that for news co-mentioning, the long-short portfolio based on within-industry and cross-industry momentum spillovers generate very similar returns and CH-4 adjusted alphas. In contrast, for other linkage types, the within-industry momentum spillover is noticeably stronger. For example, in the case of geographic and technology linkages, we nd that only the long-short portfolio based on the within-industry momentum spillover e ect generates a positive and statistically signi cant alpha. The cross-industry momentum spillover e ect does not yield signi cant abnormal returns. In the case of analyst co-coverage links, we nd that although its cross-industry momentum generates a positive alpha, it is marginally signi cant and only half the size of that obtained from the within-industry momentum. This exercise provides evidence that news co-mentioning reveals a wider range of economic links, and news-implied linkages are more comprehensive compared to other types of linkage proxies. Speci cally, news co-mentioning incorporates both industry information and valuable non-industry linkage information, which contributes to its strong predictive power for future returns.

In addition to containing comprehensive information, another advantage of news is its prompt update, which helps us to identify changes in linkages among rms in time. We keep track of the percentage change in di erent linkage networks over time, and we nd that the news-implied network updates faster than other networks. Every week, some links are added to the new network, and some stale links are removed. Due to the fast update, it is more challenging for investors to gather and process this information, which might explain why news co-mention momentum spillover is stronger. In contrast, other networks are more persistent, making it relatively easier for investors to take this information into account when making decisions.

After con rming that news co-mention momentum spillover uni es various momentum spillover e ects and highlighting the advantages of news in identifying linkages, we further investigate the mechanism behind this news-based momentum spillover. We nd that the news co-mention momentum is weaker among rms with higher analyst coverage, more analyst reports, larger oat values, higher institutional holder proportions, less opacity, and less linkage complexity. These ndings provide support for the limited attention explanation of news co-mention momentum.⁵ Recent work by Huang et al. (2021) proposes a new behavioral explanation for

(2023), and others).

The paper relates to the literature that extracts soft information from textual data. In recent years, there has been an explosion of empirical research in economics and nance that utilizes text as data. To mention a few examples, Ke et al. (2019) utilize information from news articles for predicting asset returns, Hu et al. (2021) extract information from Reddit to study the impact of social media on market dynamics, and Cong et al. (2019) develop a framework to generate 'textual factors' from large text datasets. News data has received signi cant attention among various alternative data sources, and there is a growing interest in using news coverage to identify linkages among rms. Studies such as Scherbina and Schlusche (2013) have demonstrated the usefulness of rm linkages identi ed through news co-mentioning in predicting stock returns. Additionally, Schwenkler and Zheng (2019) have developed a machine-learning method to construct a news-implied network of rms, and they also applied the same identi cation strategy to establish crypto peers (Schwenkler and Zheng (2021)). Guo et al. (2017) have explored the association between news co-mentioning and investor attention spillovers. This paper adopts two di erent identi cation strategies to infer rm linkages from news, and we show that sentence co-mentioning strategy is more e ective in identifying genuine links among rms. Unlike the ndings of Scherbina and Schlusche (2013), who use article co-mentioning to identify linkages and found that trading strategies based on the predictability of linked rms' past returns may not be pro table when considering transaction costs, we demonstrate that our strategy, based on sentence co-mentioning to infer links, remains robust and pro table even when accounting for real-world transaction costs. In addition, we show that due to the unique advantages of news information, all di erent forms of cross- rm momentum e ects can be uni ed by the news co-mention momentum spillover in the Chinese stock market.

The rest of the paper is organized as follows. In Section 2, we introduce the data and the construction of di erent linkages. Section 3 conducts the portfolio analysis for each type of momentum spillover e ect. In Section 4, we examine the unifying e ect of the news co-mention momentum spillover and explore the di erences in the information content of links implied by news co-mentioning and other proxies. Section 5 investigates the mechanism underlying the news co-mention momentum. In Section 6, we conduct some further analysis and robustness checks. Section 7 concludes. Additional materials are given in Appendix.

2 Data and summary statistics

The sample stocks used in this paper include all A-shares listed on the main boards of the Shanghai Stock Exchange (SSE), Shenzhen Stock Exchange (SZSE), and the Growth Enterprise Market (GEM). Special treatment (ST) shares are excluded. In the robustness analysis, we also exclude rms that have the bottom 30% capitalization from the sample. This is done to prevent potential biases related to the shell e ect in the Chinese stock market, as highlighted in the study by Liu et al. (2019). The full sample period for our analysis spans from 2006 to 2020. Due to data availability issues speci c to each linkage type, the sample periods may vary and be shorter when conducting analysis for di erent types of linkages. Stock trading data, nancial statements, and risk-free interest rates (one-year deposit rate) are from CSMAR. The CH-4 factor data is obtained from Robert

⁶Since April 22, 1998, the Shanghai and Shenzhen stock exchanges have implemented Special Treatment (ST) for the stock trading of listed rms with abnormal nancial conditions. ST shares are pre xed with "ST" and are considered to have extremely high risk. Therefore, they are often excluded from research about the Chinese stock market.

F. Stambaugh's website?

In later subsections, we provide details on the data sources and construction methodology for each type of economic linkage considered in the study, as well as the corresponding predictive signals derived from these linkages. We rst discuss the news co-mention links, which are our main focus. Then, we provide brief descriptions of seven other proxies for inter- rm linkages commonly studied in the literature, including the analyst co-coverage (Ali and Hirshleifer, 2020), customer-supplier relationship (Cohen and Frazzini, 2008), industry classi cation(Hou, 2007), geographic proximity (Parsons et al., 2020), technology a nity (Lee et al., 2019), statistical similarity of rm features (He et al., 2021), and concept category (Du et al., 2022).

2.1 News Co-mention Linkages

We use millions of news articles from the Financial Text Intelligent Analysis Platform of RESSET and the Juyuan Database spanning from 2006 to 202[®]. We Itered out 1,138,247 news articles that mentioned at least one listed rm in the A market. Table 14 from Appendix A presents a summary of the daily basic information of the news data since 2006. Prior to 2012, the news data is relatively sparse, with an average of fewer than 100 news pieces per day. The news data has become much more abundant since 2012. Given that, although the news data has been available since 2006, in the main body of the paper, we use the subset from 2012 to 2020 when the news data quality is high. For robustness check, we also consider di erent sample periods.

We adopt two strategies to construct the news co-mention linkages. The rst identi cation strategy is referred to assame article, which de nes two rms as linked if they are co-mentioned in the same piece of news article during a pre-speci ed identi cation window. One potential issue with this identi cation strategy is that two rms that happen to appear in the same article may appear in di erent sentences and are actually unrelated to each other. Given this concern, the second identi cation strategy we adopt is called same sentence which de nes two rms as linked if they are co-mentioned in the same sentence of the same piece of news article during a pre-speci ed identi cation window.

For both identi cation strategies, we consider identi cation windows of various lengths, including 3 months, 6 months, 9 months, and 12 months. The identi cation window cannot be too short or too long in order to strike a balance between capturing an adequate amount of news co-mentions and the inclusion of very stale information. Short windows (e.g., 1-month) result in a small number of stocks with news links, while long windows (e.g., over one year) include outdated news with limited predictability. At the end of each week, for a focal rm i, we calculate the news-based predictive signal as the average excess return of linked rms weighted by the number of co-mentions during the identi cation window. To be precise, we have eight news-based signals with di erent identi cation strategies and window lengths. For example, the signal constructed based on same sentence and 3-month window is denoted asSentence3_Rtn, and the corresponding trading strategy is referred to asSentence3.

⁷The website provides monthly and daily CH-4 factor data. We construct weekly factors by using cumulative daily factor returns in one week (i.e., at the end of each trading week, the weekly CH-4 factor is the cumulative return of daily CH-4 factor returns from **2**(e3564(rst)-3w)3085(2(-354(rst)-3w)30t0 0 m 3259£8574) eigtn)-3wfactor7

Table 1: Summary statistics

Link type	Variables	Count	Mean	Std.	Min.	Median	Max.
	# Stocks	464	1100	280	715	996	1840
	# Peer rms	510429	9	16	1	3	235
Sentence_3	Value	510429	22.45	71.08	0.27	7.19	2509.88
	Peer rm return	510429	0.0050	0.0697	-0.4101	0.0034	24.1232
	# Stocks	464	1515	231	1164	1462	2324
	# Peer rms	702802	12	21	1	4	311
Sentence_6	Value	702802	19 12	61.36	0.27	6 76	2509.88
	Peer rm return	702802	0 0044	0.0641	-0.4102	0.0036	24 1232
	# Stocks	464	1758	187	1316	1756	2502
	# Peer rms	815530	15	25	1	5	321
Sentence_9	Value	815530	17 73	57 27	0.27	6 56	2509.88
	Peer rm return	815530	0.0038	0.0555	-0.4101	0.0036	1 5622
	# Stocks	464	1925	199	1437	1915	2613
	# Peer rms	803012	17	28	1	6	355
Sentence_12	//alue	893012	16.91	20 54 89	0.27	641	2509 88
	Poor rm roturn	803012	0.0034	0.0540	-0.4103	0.41	1 07/1
	# Stocks	464	1332	203	903	123/	2223
	# Door rmc	619272	16	233	1	5	151
Article _3		618272	20.41	65 11	0.27	5	2500.88
	Value Door rm roturn	619272	20.41	0.0550	0.27	0.90	2009.00
	# Stocko	464	1750	0.0009	-0.4101	1742	3.3024
	# Deer me	404	1750	233	1555	1742	2704
Article _6	# Peer Ins	012000	23	42	0.07	0	031
		012000	17.79	57.39	0.27	0.00	2009.00
	Peer rm return	812053	0.0041	0.0526	-0.4101	0.0044	0.8516
	# Stocks	464	1976	229	1478	1952	2816
Article _9	# Peer rms	917064	29	51	1	10	/5/
	Value	917064	16.69	54.22	0.27	6.38	2509.88
	Peer rm return	917064	0.0036	0.0507	-0.4097	0.0042	1.0741
	# Stocks	464	2122	278	1569	2121	2891
Article 12	# Peer rms	984555	35	59	1	12	867
	Value	984555	16.06	52.43	0.27	6.24	2509.88
	Peer rm return	984555	0.0033	0.0493	-0.4103	0.0041	1.0741
	# Stocks	768	1326	348	476	1429	1872
Analyst	# Peer rms	1018102	98	84	1	75	609
, indigot	Value	1018102	15.45	57.09	0.18	6.60	6285.82
	Peer rm return	1018102	0.0032	0.0445	-0.4099	0.0049	0.9588
	# Stocks	566	109	56	24	71	197
Customer	# Peer rms	61428	1	0	1	1	5
Oustonier	Value	61428	8.37	20.68	0.08	4.53	517.98
	Peer rm return	61428	0.0025	0.0580	-0.3662	0.0008	2.6103
	# Stocks	768	2336	795	1048	2313	3893
Industry	# Peer rms	1794095	130	83	2	110	364
maastry	Value	1794095	13.85	59.28	0.11	4.88	6285.82
	Peer rm return	1794095	0.0039	0.0463	-0.2713	0.0058	0.5055
	# Stocks	768	2355	783	1088	2329	3892
Geographic	# Peer rms	1808617	201	162	3	157	651
Ocographic	Value	1808617	13.84	59.05	0.09	4.88	6285.82
	Peer rm return	1808617	0.0040	0.0445	-0.2885	0.0063	1.0308
	# Stocks	768	1020	513	318	954	2117
Technology	# Peer rms	783093	680	501	1	599	2109
recimology	Value	783093	17.99	82.98	0.08	5.27	6285.82
	Peer rm return	783093	0.0040	0.0437	-0.3096	0.0064	0.9269
	# Stocks	768	2015	704	988	2014	3548
Statistical	# Peer rms	1547320	263	79	115	240	390
Statistical	Value	1547320	14.19	56.96	0.09	5.03	2619.90
	Peer rm return	1547320	0.0032	0.0453	-0.2879	0.0052	0.2410
	# Stocks	227	2970	497	2044	3028	3821
Concert	# Peer rms	674244	433	504	1	260	2704
Concept	Value	674244	17.14	60.78	0.48	6.21	2509.88
	Peer rm return	674244	0.0005	0.0340	-0.1711	0.0006	0.6113

This table reports summary statistics. The sample stocks include all listed stocks on the main board of the Shanghai Stock Exchange, Shenzhen Stock Exchange, and Growth Enterprise Market (GEM). ST shares are excluded. The news co-mention linkages are based on either same_sentence or same_article strategy, using identi cation windows of 3 months, 6 months, 9 months, and 12 months. The industry linkage is based on the Shenwan-1 classi cation. The geographic linkage is at the province level. The sample period of news co-mention momentum is 2012-2020. The sample period of customer-supplier momentum is 2010-2020. The sample period of concept momentum is Aug. 2016-2020. The sample periods for other momentums are 2006-2020.

2.2 Other Linkages

Next, we will describe the construction of other types of linkages in order to conduct a comprehensive study on momentum spillovers in the Chinese stock market.

2.2.1 Analyst Co-coverage Linkages

The analyst reports data used in our study are from the Chinese Research Data Services Platform (CNRDS). This platform functions similarly to Wharton Research Data Services (WRDS) and provides access to a wide range of Chinese research data in the elds of nance and economics. The sample period for our study is from January 2005 to December 2020, which covers the earliest available analyst prediction data on the CNRDS platform. There is a total of 1,478,413 pieces of analyst predictions over the sample period. After removing data with null values and eliminating duplicates based on analyst codes, report dates, and stock codes, we are left with 530,696 unique pieces of analyst prediction data for the identi cation of shared-analyst linkage.

We adopt the approach outlined by Ali and Hirshleifer (2020) to identify shared-analyst linkages and compute the shared-analyst peer rm return, denoted as Analyst _Rtn . At the end of each trading week, two rms are de ned as the shared-analyst peer rms if they are co-covered by the same analyst teams in the past 12 months.⁹ At the end of each week, for a focal rm i, the shared-analyst peer rm return, Analyst _Rtn _{it} , is the average excess return of analyst peer rms, weighted by the number of shared analyst teams. Since a one-year identi cation window is required to identify the shared-analyst linkages, the predictive signals can be computed from the beginning of 2006.

2.2.2 Industry Linkages

There are three frequently used industry classi cation systems in China, namely CSRC, CITIC, and Shenwan.¹⁰ The historical information about the industry classi cations used in the study is obtained from the RESSET database, where we collect yearly data for each classi cation level of CSRC, CITIC, and Shenwan. We mainly focus on the primary classi cation of Shenwan as it is the most popular classi cation system in the Chinese nancial sector. At the end of each week, for a focal rmi, the industry peer rm return, denoted as Industry _Rtn_{it}, is calculated as the equal-weighted average excess return of all other stocks from the same industry.

2.2.3 Customer-supplier Linakges

We follow Cohen and Frazzini (2008) and extract the information about the top-5 customers for all listed rms from their annual reports. However, for Chinese listed rms, the disclosure of the real names of their top customers is not mandated by China Securities Regulatory Commission (CSRC). Due to the voluntary nature of the disclosure, many rms choose to use digits, English letters, or Chinese numbers to represent their customers

⁹In the Chinese stock market, analyst reports are often produced by a team of analysts rather than individual analysts. In our study, we consider all analysts within a team as a cohesive unit and do not distinguish them as individual analysts. In other words, two stocks are de ned as shared-analyst peer rms only if they are covered by the same analyst team. This approach acknowledges that analyst reports are collaborative e orts that incorporate the ideas and insights of the entire team.

¹⁰ In 2020, for the CSRC system, there are 19 primary industries and 81 secondary industries in the sample. For the CITIC system, there are 29 primary industries, 83 secondary industries, and 188 tertiary industries in the sample. For the Shenwan system, there are 28 primary industries, 104 secondary industries, and 227 tertiary industries in the sample.

instead of providing their real names or stock codes. This makes it challenging to identify rms' customers using the dataset. Moreover, some disclosed customers are not listed rms (they can be local governments, schools, and individuals). As a result, the number of customer-supplier linkages identified is much smaller than other linkages. Summary statistics from Table 1

closest Euclidean distance on ve rm-level characteristics: closed stock price (*P*), rm size (*SIZE*), book-tomarket ratio (*BM*), ROE (*ROE*), and assets growth (*AG*).¹³ The distance between rm *i* and rm *j* is be computed as:

$$d_{ij} = \bigcap^{(1)} \frac{1}{(P_i - P_j)^2 + (SIZE_i - SIZE_j)^2 + (BM_i - BM_j)^2 + (ROE_i - ROE_j^2) + (AG_i - AG_j)^2}.$$

At the end of each week, for a focal rm i, we calculate the distances of all other rms with it, and the top 10% of rms with the smallest distances are considered as the statistical peers of rm i.¹⁴ The statistical peer rm return *Statistical_Rtn_{it}* is then calculated as the equal-weighted excess return of its statistical peer rms.

2.2.7 Concept Linakges

In the stock market, one \concept" refers to a group of stocks that share a speci c trend or topic, such as new energy or e-commerce. Following Du et al. (2022), two rms are considered concept peers if they belong to the same concept. The historical records of concepts and constituent stocks used in our study are obtained from the publicly accessible RESSET database.¹⁵ Due to data availability, the sample period for concept momentum analysis in our study spans from August 2016 to December 2020. During this period, we identify a total of 336 concepts. Table 1 shows that for each week during the sample period, there are 2,970 sample stocks with concept linkages on average. At the end of each week, for a focal rm *i*, the concept peer rm return $Concept_Rtn_{it}$ is the weighted average returns of its concept peers, where the weight is the number of common concepts.

2.3 Summary Statistics

Table 1 shows the summary statistics for di erent linkage types. Under *Sentence_3*, each focal rm has nine peer rms on average, fewer than 16 peers from *Article_3*. This aligns with our expectations as the *same_sentence* strategy removes the potential noise links from the *same_article* strategy, resulting in fewer links identi ed. Furthermore, the number of peer rms identi ed increases with the length of the identi cation window. For

3 Portfolio Analysis

In this section, we examine the relationship between past peer rm return signals computed using various types of linkages and future stock returns. At the end of each trading week, all sample stocks are sorted into quintiles according to peer rm return signals computed using various linkages. Stocks are equal-weighted within each quintile group. The long-short portfolio involves buying the highest group and selling the lowest group. All portfolios are held for one week and are rebalanced weekly. In addition to reporting the average return over the risk-free rate, we also calculate the alpha using the CH-4-factor model. The weekly returns and alphas are then converted to a monthly frequency for better comparability with existing literature.

3.1 News Co-mention Momentum Spillover

We rst report the main results for news co-mention momentum spillovers. Table 2 presents the results of eight trading strategies based on di erent identi cation strategies (*same_sentence* and *same_article*) and window lengths (3 months, 6 months, 9 months, and 12 months). Panel A and Panel B show the excess returns, and CH-4 adjusted alphas of each portfolio, respectively.

Overall, the eight long-short portfolios that exploit news co-mention momentum spillovers exhibit statistically and economically signi cant positive excess returns, which can not be explained by the CH-4 factors, including the market, size, value, and abnormal turnover rate. Among the eight strategies, *Sentence_3* performs the best and yields a monthly return of 1.94% (t-statistic=5.33), and even the worst performing one (*Article_12*) yields a monthly return of 1.36% (t-statistic=4.49). The Spearman correlation coe cients between the rank of news-based peer rm returns and the long-only portfolio returns are equal to one or close to one. This indicates a positive and monotonic relationship between the two variables, suggesting that higher news-based peer rm returns are associated with higher long-only portfolio returns.

When comparing the eight news-based trading strategies that employ di erent identi cation strategies and identi cation window lengths, we observe that while the overall news co-mention momentum spillover e ect is strong, there are substantial di erences among the eight combinations. First, give the same identi cation window length, *same_sentence* strategy always performs better than *same_article* strategy. Taking the 3-month windows as an example, *Sentence_3* strategy generates a long-short return and alpha of 1.94% (t-statistic=5.33) and 1.86% (t-statistic=5.23) respectively, higher than 1.63% (t-statistic=1.63) and 1.58% (t-statistic=5.16) for that of *Article_3*. The results are similar for other identi cation windows, indicating that *same_sentence* might be more e ective in identifying genuine links among rms. This is because when two rms appear in the same article but in di erent sentences, there are chances that they are unrelated to each other.

Panel A: Excess returns								
		same_s	sentence			same	article	
Identi cation windows	3-month	6-month	9-month	12-month	3-month	6-month	9-month	12-month
1 (Low)	0.74	0.81	0.85	0.84	0.89	0.88	0.85	0.84
	(0.88)	(0.96)	(1.02)	(1.00)	(1.07)	(1.06)	(1.02)	(1.01)
2	1.06	1.06	1.01	0.94	1.08	1.06	1.01	1.00
	(1.32)	(1.31)	(1.23)	(1.15)	(1.32)	(1.29)	(1.23)	(1.21)
3	1.11	1.16	1.09	1.15	1.20	1.15	1.12	1.09
	(1.38)	(1.42)	(1.34)	(1.38)	(1.49)	(1.40)	(1.35)	(1.31)
4	1.38	1.44	1.38	1.37	1.33	1.40	1.39	1.41
	(1.65)	(1.71)	(1.65)	(1.63)	(1.60)	(1.67)	(1.68)	(1.68)
5 (High)	2.68	2.38	2.30	2.22	2.54	2.34	2.26	2.21
	(2.96)	(2.66)	(2.57)	(2.48)	(2.84)	(2.66)	(2.52)	(2.49)
5-1	1.94	1.55	1.44	1.37	1.63	1.45	1.40	1.36
	(5.33)	(4.97)	(4.68)	(4.49)	(5.30)	(5.14)	(4.62)	(4.49)
SpearmanR	1.00	1.00	1.00	1.00	1.00	1.00	1.00	1.00
P-value	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00

Table 2: The news co-mention momentum spillover: portfolio sorting results

Panel B: CH-4 adjusted alphas

		same_s	sentence			same.	article	
Identi cation windows	3-month	6-month	9-month	12-month	3-month	6-month	9-month	12-month
1 (Low)	0.41	0.48	0.51	0.49	0.55	0.53	0.51	0.50
	(0.49)	(0.56)	(0.61)	(0.59)	(0.66)	(0.63)	(0.61)	(0.60)
2	0.78	0.74	0.69	0.61	0.77	0.74	0.67	0.66
	(0.98)	(0.92)	(0.84)	(0.75)	(0.95)	(0.91)	(0.82)	(0.80)
3	0.77	0.81	0.71	0.76	0.89	0.81	0.77	0.73
	(0.98)	(1.01)	(0.88)	(0.94)	(1.11)	(1.00)	(0.94)	(0.88)
4	1.06	1.10	1.06	1.04	0.99	1.06	1.04	1.07
	(1.29)	(1.35)	(1.30)	(1.27)	(1.21)	(1.30)	(1.28)	(1.29)
5 (High)	2.27	2.00	1.92	1.84	2.14	1.95	1.87	1.83
	(2.66)	(2.38)	(2.26)	(2.17)	(2.52)	(2.34)	(2.21)	(2.16)
5-1	1.86	1.51	1.41	1.35	1.58	1.42	1.35	1.32
	(5.23)	(4.89)	(4.59)	(4.38)	(5.16)	(4.98)	(4.45)	(4.35)
SpearmanR	0.90	1.00	1.00	1.00	1.00	1.00	1.00	1.00
P-value	0.04	0.00	0.00	0.00	0.00	0.00	0.00	0.00

This table reports the portfolio sorting results of the news co-mention momentum spillover e ect of the *same_sentence* type and *same_article* type under the four identi cation windows, respectively. *Same_sentence* strategy considers rms appearing in the same sentence of the same news article during a given identi cation window as news co-mention peers, while the *same_article* strategy considers rms appearing in the same article during a given identi cation window as news co-mention peers. The identi cation windows include 3-month, 6-month, 9-month, and 12-month. The news co-mention peer rm return of a focal rm is computed as the average excess return of news co-mention peer rms of the focal rm during the identi cation window, weighted by the number of co-mentions. At the end of each trading week, all sample stocks are sorted quintiles based on the news co-mention peer rm returns. Within each quintile group, the stocks are equally weighted. The long-short portfolio involves buying the highest group and selling the lowest group. All portfolios are held for one week and are rebalanced weekly. The sample period is 2012-2020. Panel A reports the excess returns of portfolios, and Panel B reports the intercepts (alphas) of the regression of the returns on CH-4 factors (market, size, value, abnormal turnover rate) (Liu et al., 2019). All weekly returns and alphas are converted into monthly percentages using compound interest. SpearmanR reports the Spearman correlation coe cient between the portfolio return and the serial number for each sorting. Newey and West (1987) adjusted t-statistics are shown in parentheses. Long-short returns/alphas with t-statistics higher than 2.00 are highlighted in bold.

In addition, given the link identi cation strategy, the returns of the long-short portfolio decrease as the length of the identi cation window increases. Unlike other linkages such as industry links and technology links, which tend to be more stable and persistent in the long term, news is highly time-sensitive and re ects recent information from the market. News-based linkages capture the immediate market sentiment, investor behaviors, and breaking news events, which can quickly change over time. As a result, the news linkages are less persistent compared to other linkages and can exhibit rapid shifts as new information becomes available. The use of a long identi cation window may include outdated news that is no longer relevant. This might explain the observed results. However, a shorter identi cation window results in a smaller number of sample stocks with news-implied linkages. For instance, if we adopt *same_sentence* strategy and use a 1-month window to identify

Panel A: Exc	cess returns								
	News co-mention	Analyst	Industry	Province	City	Customer	Technology	Statistical	Concept
1 (Low)	0.74	1.18	1.35	1.51	1.61	0.92	1.54	0.91	-0.45
	(0.88)	(1.59)	(1.82)	(2.06)	(2.15)	(1.19)	(2.05)	(1.22)	(-0.50)
2	1.06	1.28	1.52	1.64	1.67	0.22	1.66	1.35	-0.15
	(1.32)	(1.80)	(2.05)	(2.21)	(2.25)	(0.27)	(2.21)	(1.75)	(-0.17)
3	1.11	1.61	1.74	1.77	1.78	0.17	1.94	1.71	0.10
	(1.38)	(2.24)	(2.28)	(2.37)	(2.41)	(0.22)	(2.56)	(2.25)	(0.11)
4	1.38	2.03	1.90	1.84	1.81	0.95	1.92	2.09	0.21
	(1.65)	(2.81)	(2.49)	(2.44)	(2.40)	(1.20)	(2.52)	(2.79)	(0.25)
5 (High)	2.68	2.04	2.12	1.86	1.79	0.38	2.12	2.28	0.36
	(2.96)	(2.74)	(2.74)	(2.42)	(2.35)	(0.48)	(2.72)	(3.03)	(0.44)
5-1	1.94	0.86	0.77	0.35	0.18	-0.53	0.58	1.36	0.81
	(5.33)	(3.33)	(3.25)	(2.77)	(2.05)	(-1.46)	(2.73)	(3.93)	(2.25)
SpearmanR	1.00	1.00	1.00	1.00	0.90	0.10	0.90	1.00	1.00
P value	0.00	0.00	0.00	0.00	0.04	0.87	0.04	0.00	0.00
Panel B: CH	-4 adjusted alphas								
	News co-mention	Analyst	Industry	Province	City	Customer	Technology	Statistical	Concept
1 (Low)	0.41	0.82	0.98	1.13	1.23	0.72	1.11	0.58	-0.40
	(0.49)	(1.09)	(1.29)	(1.52)	(1.61)	(0.94)	(1.45)	(0.75)	(-0.44)
2	0.78	0.92	1.12	1.24	1.31	0.16	1.26	0.95	-0.11
	(0.98)	(1.26)	(1.48)	(1.64)	(1.74)	(0.21)	(1.65)	(1.22)	(-0.13)
3	0.77	1.29	1.32	1.39	1.37	0.15	1.61	1.29	0.10
	(0.98)	(1.75)	(1.70)	(1.82)	(1.80)	(0.19)	(2.10)	(1.67)	(0.11)
4	1.06	1.72	1.48	1.42	1.43	0.67	1.54	1.65	0.19
	(1.29)	(2.38)	(1.94)	(1.84)	(1.87)	(0.87)	(1.97)	(2.17)	(0.21)
5 (High)	2.27	1.68	1.74	1.48	1.36	0.12	1.71	1.93	0.33
	(2.66)	(2.21)	(2.20)	(1.89)	(1.75)	(0.16)	(2.14)	(2.51)	(0.39)
5-1	1.86	0.85	0.76	0.34	0.13	-0.60	0.60	1.34	0.74
	(5.23)	(2.97)	(3.00)	(2.56)	(1.37)	(-1.43)	(2.64)	(3.35)	(1.89)
SpearmanR	0.90	0.90	1.00	1.00	0.70	-0.70	0.90	1.00	1.00
P value	0.04	0.04	0.00	0.00	0.19	0.19	0.04	0.00	0.00

Table 3: Other momentum spillovers: portfolio sorting results

This table reports the portfolio sorting performance of the shared-analyst momentum, industry momentum, geographic momentum, customer momentum, technology momentum, statistical momentum, and concept momentum. For comparison, the result for the news co-mention momentum of *same_sentence* type under the 3-month identi cation window (*Sentence_3*) is shown in the rst column. The industry momentum spillover is based on the rst level of Shenwan classi cation (Shenwan-1), and the results are shown in column 2. The geographic momentum spillover is constructed at both the province and city levels, and the results are shown in columns 3 and 4, respectively. Columns 5-8 present the results for customer momentum, technology momentum, statistical momentum, and concept momentum, respectively. Due to data quality, the sample period of news co-mention momentum is 2012-2020, the sample period of customer momentum is 2010-2020, the sample period of concept momentum is Aug. 2016-2020, and the sample periods for other momentums are 2006-2020. For each linkage type, at the end of each trading week, we rst calculate the peer returns of all focal rms and then sort all stocks into quintiles based on the peer rm return. Within each quintile group, the stocks are equally weighted. The long-short portfolio involves buying the highest group and selling the lowest group. All portfolios are held for one week and are rebalanced weekly. Panel A gives the excess returns of portfolios, and Panel B presents the intercepts of the regression of the returns on CH-4 factors (Liu et al., 2019) (market, size, value, abnormal turnover rate). All weekly returns and alphas are converted into monthly percentages using compound interest. SpearmanR reports the Spearman correlation coe cient between the portfolio return and the serial number for each sorting. Newey and West (1987) adjusted t-statistics are shown in parentheses. Long-short returns/alphas with t-statistics higher than 2.00 are highlighted in bold. In the Chinese stock market, the geographic momentum is not strong. Province-level momentum is statistically signi cant but the long-short strategy based on that yields only a small average return of 0.35% per month. City-level momentum is even weaker and is not signi cance after controlling for the CH-4 factors.

Overall, the news co-mention momentum spillover e ect is stronger than other types of momentum spillovers in the Chinese stock market. A natural question then arises: whether the news co-mention momentum spillover encompasses and uni es the various other momentum spillover e ects.

4 The Unifying E ect

In this section, we examine whether the news co-mention momentum spillover e ect encompasses and uni es the various other momentum spillover e ects observed in the Chinese stock market. To do so, we follow Ali and Hirshleifer (2020) to use spanning tests and Fama and MacBeth (1973) regressions. In addition, we investigate the information content di erences between news co-mentioning and other proxies to gain further insights into the features of news-implied linkages.

4.1 Factor Spanning Tests

To test whether the anomaly returns of di erent types of momentum spillovers can be explained by each other, we construct an augmented CH-4 model as follows:

$$R_t = \alpha + \beta_{MKT}MKT_t + \beta_{SMB}SMB_t + \beta_{HML}HML_t + \beta_{PMO}PMO_t + \beta_{MS}MS_t,$$

where R_t is the anomaly return of one momentum spillover factor to be explained, and the MS (momentum spillover) factor on the right-hand side is another momentum spillover factor used as an additional explanatory variable. We use \CH-4+a speci c MS factor" to indicate the set of explanatory variables used to explain the target momentum spillover factor R_t . For example, CH-4+Sentence_3 means the set of explanatory variables includes CH-4 factors plus a news co-mention momentum factor of the *Sentence_3* type. If the news co-mention momentum factor does explain the momentum spillover factor R_t , then we should observe the alpha of the model becoming smaller and potentially insigni cant after controlling for the news co-mention momentum factor. Additionally, the β_{MS} should be signi cantly positive.

Panel A of Table 4 shows the alpha of regressing the time series of the target momentum spillover factor R_t on the CH-4 factors plus a speci c *MS* factor. The column name corresponds to the name of the target factor R_t , and the row name indicates the set of variables in the augmented CH-4 model. All news co-mention momentum factors are constructed using the *same_sentence* identi cation strategy, which has shown superior performance. We also present results obtained using the *same_article* method in Table 20 in the robusteness part, which yields similar outcomes. We nd that most of the news co-mention momentum spillover factors, particularly those using a short identi cation window to establish links, are capable of explaining other momentum spillover factors. For instance, after controlling for the Sentence_3 factor, the alphas of all other momentum factor is as large as 1.34% (t-statistic=3.35), its CH-4+Sentence_3 alpha becomes only 0.17% (t-statistic=0.43). In the case of the analyst, industry, technology, and concept momentum, alphas even turn negative after the addition of the Sentence_3 factor. The results are similar for the Sentence_6, Sentence_9, and Sentence_12 factors, which all have strong explanatory power for other factors. One special case is the geographic momentum factor. Although it has insigni cant alpha in the CH-4+Sentence_3 model, the factor still has positive and signi cant alpha if one controls any of the three longer-term news co-mention factors. This is in line with our previous ndings that the strength of news co-mention momentum decreases as the length of the identi cation window increases. Actually, the Sentence_3 momentum factor can explain the news co-mention momentum factor with longer identi cation windows, but not vice versa.

Panel B of Table 4 reports the factor loadings of other momentum spillover factors on the CH-4+Sentence_3 model. All the loading coe cients of the Sentence_3 factor are signi cantly positive, and much bigger than the loadings on CH-4, indicating that all other cross-asset momentum anomalies are explained by their loadings on the Sentence_3 news co-mention momentum factor.

As a comparison, we also attempt to explain the news co-mention momentum spillover factors with other cross-momentum factors including the factors based on analyst co-coverage momentum, industry momentum, geographic momentum, technology momentum, statistical momentum, and concept momentum. We nd that none of the four news co-mention momentum spillover factors can be explained by those other momentum spillover factors as the alphas all remain large and signi cant. For example, the momentum spillover via statistical linkages yields a high monthly long-short return and CH-4 alpha, which is the second-highest among all trading strategies. However, the Sentence_3 momentum factor still generates an alpha of 1.16% (t-statistic=4.17) per month under the CH-4+Statistical model. Moreover, even the weakest news co-mention momentum spillover factor (i.e., the Sentence_12 factor) yields a CH-4+Statistical alpha of 0.51% (t-statistic=2.57) per month. The analyst co-coverage momentum factor, which was found to explain all other cross-momentum factors in the US market, cannot explain either the news co-mention momentum spillover factors or the statistical momentum factor. In the last row of panel A, we include all the six momentum spillover factors that are not related to news as additional explanatory variables. We nd that the CH-4+Non_news model is unable to explain the news co-mention momentum spillover factors. This suggests that even when we include all other non-news related momentum spillover factors as additional explanatory variables, they cannot explain the abnormal returns associated with the news co-mention momentum anomaly.¹⁷

We conduct several robustness checks, and the detailed results can be found in subsubsection 6.3.3. Firstly, we run regressions with the Sentence_3 factor as the sole explanatory variable (CH-4 factors are excluded), which is shown in Table 17. Besides, we construct industry momentum factors based on other industry classi cations, and the results are shown in Table 18. Furthermore, we perform spanning tests for di erent sub-sample periods, including the periods 2006-2020 and 2012-2020. The results are summarized in Table 19. Finally, in Table 20, we convert the co-mention momentum from *same_sentence* type to *same_article* type under the four identi cation windows. Overall, our main is not provide the test of the sentence in the sentence is changes that we made.

¹⁷In the CH-4+Non_news model, we include all six non-news factors as explanatory variables. However, we recognize the concern that the small sample size of concept momentum could potentially impact our results. To address this concern, we perform a robustness check by excluding the concept momentum variable from the model. Our results are robust to the exclusion of the concept momentum variable.

Panel A: Alphas of fact	or spanning tests									
	Sentence_3	Sentence_6	Sentence_9	Sentence_12	Analyst	Industry	Geographic	Technology	Statistical	Concept
CH-4 + Sentence_3		-0.15	-0.22	-0.30	-0.44	-0.53	0.25	-0.26	0.17	-0.51
		(-1.51)	(-1.76)	(-2.15)	(-1.76)	(-2.10)	(1.92)	(-1.11)	(0.43)	(-1.59)
CH-4+Sentence_6	0.37		-0.08	-0.16	-0.34	-0.45	0.27	-0.21	0.26	-0.41
	(3.48)		(-1.17)	(-1.80)	(-1.47)	(-1.93)	(2.11)	(-0.95)	(0.71)	(-1.37)
CH-4 + Sentence_9	0.54	0.16		-0.08	-0.25	-0.38	0.26	-0.17	0.30	-0.31
	(3.95)	(2.16)		(-1.62)	(-1.13)	(-1.76)	(2.08)	(-0.79)	(0.91)	(-1.10)
CH-4+Sentence_12	0.65	0.27	0.11		-0.17	-0.32	0.27	-0.13	0.36	-0.26
	(4.27)	(2.88)	(2.30)		(-0.81)	(-1.53)	(2.15)	(-0.60)	(1.13)	(-0.94)
CH-4+Analyst	1.18	0.80	0.67	0.58		0.11	0.24	0.17	0.70	0.06
,	(4.97)	(4.25)	(3.79)	(3.47)		(0.76)	(1.83)	(0.99)	(2.18)	(0.28)
CH-4+Industry	1.35	0.98	0.84	0.76	0.15		0.25	0.14	0.78	0.33
'n	(2.08)	(4.49)	(4.30)	(4.03)	(0.96)		(1.92)	(0.87)	(2.25)	(1.67)
CH-4+Geographic	1.62	1.28	1.13	1.06	0.59	0.55		0.44	1.05	0.34
-	(4.94)	(4.38)	(3.94)	(3.72)	(2.21)	(2.26)		(2.07)	(2.61)	(0.92)
CH-4+Technology	1.49	1.12	0.99	0.91	0.34	0.27	0.27		0.93	0.25
	(4.63)	(4.07)	(3.86)	(3.63)	(1.59)	(1.41)	(2.05)		(2.50)	(0.98)
CH-4+Statistical	1.16	0.77	0.61	0.51	0.20	0.23	0.23	0.25		0.04
	(4.17)	(3.28)	(2.90)	(2.57)	(0.90)	(0.97)	(1.69)	(1.24)		(0.14)
CH-4+Concept	1.33	0.89	0.71	0.60	0.29	-0.16	0.35	0.04	0.92	
-	(4.06)	(3.65)	(3.08)	(2.81)	(1.30)	(-0.89)	(2.19)	(0.18)	(2.61)	
CH-4+Non_news	1.15	0.72	0.52	0.40			•		•	
	(3.59)	(3.00)	(2.22)	(1.92)						
Panel B: CH-4+Sentend	ce_3 factor loading									
	Sentence_3	Sentence_6	Sentence_9	Sentence_12	Anayst	Industry	Geographic	Technology	Statistical	Concept
Sentence_3		0.896	0.878	0.886	0.835	0.748	0.086	0.518	0.862	0.747
		(28.82)	(23.66)	(21.14)	(11.93)	(10.92)	(4.09)	(7.12)	(8.50)	(6.68)
mktrf		-0.012	-0.014	-0.028	-0.085	-0.069	0.011	-0.034	-0.078	-0.012
		(-1.10)	(-0.82)	(-1.54)	(-3.55)	(-2.60)	(0.84)	(-1.42)	(-1.73)	(-0.30)
NMG		-0.008	-0.008	-0.007	-0.001	-0.018	-0.027	-0.051	0.019	0.004
		(-0.36)	(-0.25)	(-0.21)	(-0.02)	(-0.26)	(-1.00)	(-0.73)	(0.17)	(0.04)
SMB		0.010	0.006	0.016	0.044	0.054	-0.001	-0.026	0.068	0.025
		(0.46)	(0.19)	(0.42)	(0.75)	(0.85)	(-0.05)	(-0.46)	(0.61)	(0.26)
PMO		-0.032	-0.037	-0.057	-0.082	-0.036	-0.018	-0.059	-0.203	0.087
		(-2.03)	(-1.97)	(-2.05)	(-2.06)	(-0.79)	(-1.16)	(-1.49)	(-2.33)	(0.97)
Adj. R2		0.881	0.824	0.797	0.525	0.460	0.063	0.275	0.355	0.351
# Õbs.		464	464	464	464	464	464	464	464	227
Panel A reports the interce	pt (or alpha) of revision of the Λ	gressing the time <i>AS</i> factor is giver	series of di erent l	ong-short momentu The column name	um returns on the ty	ne CH-4 factors	(<i>MKT</i> , <i>SMB</i> , <i>HM</i> † momentum return	<i>L, PMO</i>) (Liu et an of the demendent	il., 2019) plus each ant variable) whil	MS cross- rm
means the augmented CH	4 model (i.e., the c	priginal CH-4 mo	del plus a speci c	MS factor). For ex	ample, the alph	ias of the row C	H-4+Sentence_3 in	dicate the alphas f	rom the time-seri	es regressions of
long-short portfolio returns	of other momentu	ims on MKT, SM	<i>IB</i> , <i>HML</i> , <i>PMO</i> ar	nd the Sentence_3 f	actor. In the lasi	t row of panel A	v, we include all the	six momentum sp	illover factors that	are not related
The leave as additional expla	hatory variables.	All news co-ment	ion momentum rac	tors are pased on s	ame_sentence []	ype under identi	Cation windows if A do not concider to	nciuaing 3-montn, 1 be etictomer memo	o-montn, y-montn striss diis to the I	, and Iz-montn.
The industry momentum is The sample period of paws	co-mention mome	ntum is 2012-202	ion, wniie the geogr '0 The sample peri	apriic momentum i	s based on the p	0016-2020 The	/e do riot corisider t samnle nerinds for	ne customer momentums	nium due to the l are 2006-2020	uw data quality.
(1987) adjusted t-statistics	are shown in pare	ntheses. Alphas v	with t-statistics high	her than 2.00 are h	iahliahted in bol	Id. Panel B repu	orts the factor load	inds of other mome	entum spillover lo	ng-short returns
on the CH-4+Sentence_3 m	odel. Coe cients	with t-statistics h	higher than 2.00 are	highlighted in bold		-		5		5
			0)						

Table 4: Factor spanning test

4.2 Fama-Macbeth Regression Tests

Our spanning tests reveal that all other cross-asset momentum anomalies can be explained by the news comention momentum spillover factor. To provide additional evidence, we employ Fama-MacBeth regressions to examine the relationship between the stock excess return in the next week and the weighted average peer rm returns constructed using di erent linkage proxies.¹⁸ Our control variables include the size (taking logarithms), book-to-market ratio, and stock return in the past week. All independent variables are standardized by their cross-sectional means and standard deviations.

We estimate two sets of regressions based on two di erent samples. The rst one takes the union of all samples of average peer rm returns constructed using di erent linkage proxies, ¹⁹ and the second sample takes the intersection set of all peer return samples. Table 5 presents the results of the two sets of Fama-MacBeth regressions. Panel A and Panel B present the results for the union sample and intersection sample, respectively. Columns 1-7 use each peer rm return alone as the dependent variable respectively. The past-one-week news co-mention peer rm return, *Sentence_3_Rtn*, exhibits the strongest predictability power for future returns, both statistically and economically. A one standard deviation increase in the past-one-week *Sentence_3_Rtn* is associated with an increase of 27.1 bps in the future return, with a t-statistic of 5.53. For the analyst and statistical momentum, which follow the news co-mention momentum according to portfolio analysis, their predictive powers are much smaller, with just 7.1 bps (t-statistic=4.32) and 3.9 bps (t-statistic=2.57) per standard deviation change of the peer rm return, respectively. The geographic peer rm return does not exhibit a signi cant predictive power of the future return based on either the union sample or the intersection sample. This result is consistent with the portfolio sorting and spanning tests.

In column 8 of Panel A, both *Sentence_3_Rtn* and *Analyst_Rtn* are included in the regression. The coe cient and t-statistic of *Sentence_3_Rtn* hardly change with the inclusion of *Analyst_Rtn*. However, the coe cient of *Analyst_Rtn* decreases substantially by 56% from 0.071 to 0.031. A one standard deviation increase in *Sentence_3_Rtn* predicts an increase of 26.5 bps in the future return, while a one standard deviation increase in *Analyst_Rtn* predicts an increase of just 3.1 bps in the future return. Telling from Panel B column 2, we reach a similar result if we use the intersection sample. These results indicate that the analyst momentum, which is found to unify all momentum spillover e ects in the US (Ali and Hirshleifer, 2020), does not possess the same predictive power in the Chinese market. Instead, its role is taken by the news co-mention momentum. When we include *Sentence_3_Rtn* and any other peer rm returns decrease signi cantly once the *Sentence_3_Rtn* is controlled. From column 9, including *Sentence_3_Rtn* largely weakens *Industry_Rtn*, whose e ect becomes insigni cant. As shown in column 5, the technology peer rm return alone is a strong predictor of future returns. However, the coe cient of *Technology_Rtn* becomes insigni cant once *Sentence_3_Rtn* is added. From columns 6 and 12, we nd that the coe cient of the statistical peer return decreases by more than 50% after the inclusion of *Sentence_3_Rtn*. The geographic and concept momentum e ects exhibit similar results.

¹⁸The news co-mention peer rm return is based on *Sentence_3* strategy; the industry peer rm return is based on the Shenwan-1 classi cation; and the geographic peer rm return is based on peers at the province level.

¹⁹We II peer rm returns of null values with 0.

Panel A: FM regr	essions bas	ed on union	samples											
Sentence_3_Rtn	1 0.271 /E E2)	2	m	4	വ	9	٢	8 0.265 /F 11)	9 0.264 /E 20)	10 0.270 /E E1)	11 0.267 /E_4E	12 0.267 /E E 2)	13 0.120 /2 14)	14 0.257 (E 20)
Analyst_Rtn	(cc.c)	0.071						0.031 0.031	(oc.c)	(10.0)	(0.4.0)	(20.0)	(01 ·c)	0.020 0.020
Industry_Rtn		(70.4)	0.040					(10.1)	0.019 (0.86)					0.007
Geographic_Rtn			(+0.2)	0.015					(00.0)	0.016				0.013 0.013
Technology_Rtn				(20.1)	0.039					(00.1)	0.028			0.019
Statistical_Rtn					(10.7)	0.039					(1:44)	0.022		0.016 0.016
Concept_Rtn						(1.04)	0.036 (1.59)					(c/.0)	0.035 (1.55)	(oc.0)
Control Ava. R Sauare	YES 0.060	YES 0.061	YES 0.064	ΥΕS 0.059	YES 0.059	YES 0.062	YES 0.046	YES 0.064	YES 0.067	ΥΕS 0.062	ΥΕS 0.062	ΥES 0.065	YES 0.048	YES 0.076
Total Obs. Avg. Obs.	1709464 2226	1709464 2226	1709464 2226	1709464 2226	1709464 2226	1709464 2226	653094 2877	1709464 2226	1709464 2226	1709464 2226	1709464 2226	1709464 2226	653094 2877	1709464 2226
Panel B: FM regr	essions base	ed on interse	ection samp	les										
Sentence_3_Rtn	0.234							0.159	0.230	0.233	0.229	0.048	0.112	0.059
Analyst_Rtn	(40.0)	0.081						(51.6) 0.026 (70.17)	(/c.c)	(60.0)	(11.6)	(11.0)	(10.0)	(2.10) 0.022 (0.65)
Industry_Rtn		(04.4)	0.041						-0.009					0.039 0.039
Geographic_Rtn			(00.2)	0.015					(0c.0-)	-0.001				0.030
Technology_Rtn				(8c.1)	0.037					(-0.11)	0.021			(1.91) 0.041 (1.04)
Statistical_Rtn					(0+.7)	0.079					(00.0)	0.100 (7 27)		0.017 0.017 0.017
Concept_Rtn							0.036 (1.59)					(12.0)	0.022 (0.74)	
Control Avg. R Square	YES 0.088	YES 0.064	YES 0.064	YES 0.059	YES 0.079	YES 0.058	YES 0.046	YES 0.086	YES 0.096	YES 0.090	YES 0.125	YES 0.070	YES 0.067	YES 0.131
Total Obs. Avg. Obs.	503829 1086	988653 1287	1694481 2206	1707853 2224	750982 978	1456040 1896	653094 2877	305452 658	501611 1081	503511 1085	232484 501	441002 950	205109 904	106708 230
This table presents is the peer rm retu	the results urn for eacl	of the Fama h type of eco	and MacB onomic link	eth (1973) r age constru	regressions. Icted in <mark>sub</mark>	The depen section 2.2.	ident varial . The cont	ble is the extronation of the extra the section of	kcess stock r ss include th	return in th he rm size	e next week (taking log	, while the arithms), b	independe ook-to-ma	nt variable irket ratio,
and stock return ir is based on <i>same_s</i> .	entence type	week. All ir	i 3-month i	variables ai denti catior	re standarc n window.	lized with t The indust	their cross- ry moment	sectional m tum is base	teans and s and the St	tandard dev henwan-1 cl	viations. T lassi cation	he news co- . The geog	-mention n raphic mo	nomentum mentum is
conducted on the p sample period of cc	rovince leve incept morr	el. We do no tentum is A	ot consider t ug. 2016-20	he customei 120. The sai	r momentui mple periou	m due to th ds for other	e low data · momentu	quality. Th ms are 2000	he sample pi 6-2020. We	eriod of new estimate th	/s co-mentic he regressio	on momentuns based or	um is 2012- two samp	2020. The cles. Panel
A uses the union si time series). Panel	et of sampl. B uses the	es, and peer intersection	r rm returi set of all sé	ns of null vi amples of di	erent pee	lled with 0 r rm retun	(concept n ns. Colum	nomentum ns 1-7 use €	samples are sach peer i	e excluded f rm return a	rom taking lone as the	the union dependent	due to its variable re	very short spectively.
n columns 8-13, S egression. Newey a	entence_3_1 and West (Rtn is addec 1987) adjust	to the regret to the regret to the test of	ressions. In ics are show	column 14 /n in paren	, all peer i theses. For	rm returns brevity, al.	except for I coe cient	<i>Concept_R</i> s are multip	tn (due to 1 plied by 100	the very sho).	ort ⁱ time ser	ies) are ad	ded to the

Table 5: Fama-MacBeth regressions

As indicated in columns 4, 10, and 7, 13 of Panel A and Panel B, the coe cients of *Geographic_Rtn* and *Concept_Rtn* are not signi cant, whether or not we control for news co-mention momentum variable in the regression. Particularly, in the case of the intersection sample, the coe cient of *Geographic_Rtn* even becomes negative after controlling for *Sentence_3_Rtn*. After all, according to the previous portfolio sorting analysis, the geographic and concept momentum e ects lack both statistical and economic signi cance.

Finally, in column 14 of Table 5, we show the FM regression results with all peer rm returns included.²⁰ *Sentence_3_Rtn* remains to be a strong predictor of the future stock return after controlling for all of the previously documented cross- rm momentums simultaneously. Based on the union sample, Panel A reveals that *Sentence_3_Rtn* exhibits the strongest predictive power among all past peer rm returns. A one standard deviation increment in *Sentence_3_Rtn* is associated with an increase of 25.7 bps in future returns (with a t-statistic of 5.29). In contrast, all other peer rm returns lose predictive power in this big regression.

Overall, the results of the Fama-MacBeth regressions are consistent with the previous portfolio sorting analysis and factor-spanning test, providing further evidence that the past news co-mention peer rm return exhibits a strong and robust predictive power for future stock returns. Moreover, the predictive power of news co-mention peer rm return largely subsumes the predictability of other types of momentum spillover e ects. This holds true for both the union sample and the intersection sample.

4.3 Further Analysis

The factor-spanning test and Fama-MacBeth regressions show that all di erent forms of cross- rm momentum e ects can be uni ed by the news co-mention momentum spillover in the Chinese stock market. In this sub-section, we delve deeper into the di erences in the information content of links implied by news co-mentioning and other proxies, aiming to gain a better understanding of the unique characteristics and e ectiveness of news co-mentioning as a linkage identi cation strategy compared to alternative methods.

As argued by Scherbina and Schlusche (2013) and Schwenkler and Zheng (2019), one of the advantages of news-implied economic linkages is their ability to reveal a broader range of economic linkages. To examine whether news co-mention momentum spillover is stronger than others because news-implied linkages are more comprehensive, we conduct an exercise to analyze the within- and cross-industry momentum spillovers separately for each type of momentum spillover in subsubsection 4.3.1. Another advantage of news is its prompt update, which helps us to identify changes in linkages among rms in time. We also conduct a small exercise in subsubsection 4.3.2 to investigate that aspect.

4.3.1 Cross- and Within-Industry Momentum Spillovers

Unlike the US market, where the shared-analyst momentum serves as the unifying momentum spillover e ect, our ndings in the Chinese stock market suggest that the momentum driven by news-co-mention linkages, rather than shared-analyst linkages, can encompass the predictability of other momentum spillover e ects. This discrepancy could be attributed to the di erences in the information content of these linkages. One characteristic of Chinese analysts is that each of them tends to cover rms in one industry. This might lead to a high degree of

²⁰Concept peer rm returns are excluded from the analysis due to its very short time series (from Aug. 2016 to 2020). Including concept returns would signi cantly shorten the sample time series.

information overlapping between the shared-analyst linkage and industry linkage.²¹ News, on the other hand, might reveal a broader range of economic linkages (i.e., a lower degree of overlapping with industry links).

											1.0
Sentence_3 -	- 1	0.94	0.91	0.89	0.71	0.67	0.23	0.52	0.58	0.6	- 1.0
Sentence_6 -	0.94	1	0.97	0.96	0.78	0.74	0.24	0.59	0.65	0.66	- 0.9
Sentence_9 -	0.91	0.97	1	0.99	0.8	0.77	0.27	0.61	0.68	0.69	- 0.8
Sentence_12 -	0.89	0.96	0.99	1	0.81	0.78	0.28	0.62	0.71	0.71	- 0.7
á mai y air	0.71	0.78	0.8	0.81	1	0.84	0.28	0.66	0.6	0.78	
Industry -	0.67	0.74	0.77	0.78	0.84	1	0.26	0.7	0.54	0.87	- 0.6
Geographic -	0.23	0.24	0.27	0.28	0.28	0.26	1	0.23	0.26	0.32	- 0.5
Technology -	0.52	0.59	0.61	0.62	0.66	0.7	0.23	1	0.42	0.78	- 0.4
Statistical -	0.58	0.65	0.68	0.71	0.6	0.54	0.26	0.42	1	0.54	
Concept -	0.6	0.66	0.69	0.71	0.78	0.87	0.32	0.78	0.54	1	- 0.3
	Sentence_3	Sentence_6	Sentence_9	Sentence_12	4 maily st	industry	Geographic	Technology	Statistical	Concept	

intra-industry peers only. By exploring both within-industry and cross-industry e ects, this paper provides a more comprehensive understanding of how information propagates among rms in the Chinese stock market.

Table 7 reports the long-short portfolio excess returns and CH-4 alphas of the news co-mention, analyst, geographic, technology, statistical, and concept momentum, separately considering cross-industry and within-industry links. For the news co-mention linkage, there is not much di erence in the predictive power of cross-industry peer returns and within-industry peer returns. The cross-industry news co-mention momentum yields an average excess return of 1.73% (t-statistic=4.95), while the within-industry news co-mention momentum yields 1.76% (t-statistic=3.92). From Panel B, the cross-industry and within-industry news co-mention momentum mentum generates the same CH-4 adjusted alphas at 1.65% (t-statistics = 5.06 and 3.74, respectively) per month.

Panel A: Excess r	return					
	News co-mention	Analyst	Geographic	Technology	Statistical	Concept
Cross Industry	1.73	0.45	0.10	0.13	0.89	0.22
	(4.95)	(2.44)	(0.95)	(0.83)	(2.92)	(0.71)
Within Industry	1.76	0.78	0.43	0.75	0.94	0.54
	(3.92)	(3.20)	(2.52)	(3.19)	(3.32)	(1.61)
Panel A: CH-4 ad	ljusted alphas					
	News co-mention	Analyst	Geographic	Technology	Statistical	Concept
Cross Industry	1.65	0.45	0.07	0.12	0.80	0.16
	(5.06)	(2.34)	(0.59)	(0.73)	(2.26)	(0.48)
Within Industry	1.65	0.75	0.42	0.79	0.89	0.43
	(3.74)	(2.82)	(2.33)	(3.08)	(2.79)	(1.19)

Table 7: Cross- and within-industry momentum spillovers: portfolio analysis

This table reports the long-short portfolio based on cross- and within-industry momentum spillovers for each linkage type. We consider ve momentum spillover e ects, including the news co-mention momentum, analyst momentum, geographic momentum, technology momentum, statistical momentum, and concept momentum. The news co-mention momentum is based on *same_sentence* type under the 3-month identi cation window. The geographic momentum is at the province level. The industry classi cation we adopt is the CSRC-1 classi cation. For the cross-industry e ect, when calculating the peer rm return, only peer rms from di erent industries from the focal rm are considered. On the contrary, for the within-industry e ect, when calculating the peer rm return, only peer rm return, respectively. The long-short portfolio involves buying the highest group and selling the lowest group. The sample period of news co-mention momentum is 2012-2020. The sample period of concept momentum is Aug. 2016-2020. The sample periods for other momentums are 2006-2020. Panel A shows the excess returns of the long-short portfolios, while Panel B reports the CH-4 adjusted alphas. All returns alphas are converted to monthly using compound interest. Newey and West (1987) adjusted t-statistics are shown in parentheses.

In contrast, for any other type of linkage, the within-industry momentum e ect is stronger. For instance, the within-industry analyst momentum generates a CH-4 alpha of 0.75% (t-statistic=2.34) per month, whereas the cross-industry analyst momentum produces a CH-4 alpha of 0.45% (t-statistic=2.34) per month. Comparing that with the alpha of the shared-analyst momentum (0.85%, with t-statistic=2.97), we nd that the predictive power mostly stems from within-industry information. We observe similar patterns for statistical linkage and concept linkage, which both have a stronger within-industry momentum e ect. For the geographic and technological linkages, only the within-industry e ects are signi cant.

Overall, these results suggest that news co-mention linkages exhibit unique characteristics compared to other types of linkages. While other types of linkages show a higher degree of overlap with the industry linkage (except for the geographic linkage), news co-mention linkages capture a broader range of economic linkages. They incorporate both industry information and valuable non-industry linkage information, which contributes to their strong predictive power for future returns. This nding helps explain why news co-mention momentum spillover e ects tend to dominate and unify other types of momentum spillovers in the Chinese stock market.

4.3.2 Speed of Update

As mentioned previously, the strong predictive power and unifying e ect of the news co-mention momentum can be attributed to two advantages of the news co-mention linkage: comprehensive information and prompt updates. subsubsection 4.3.1 demonstrates that the news content is all-encompassing, and it incorporates both industry information and valuable non-industry linkage information. From the portfolio analysis, we have seen that the predictive power of the news co-mention momentum decreases as the identi cation window gets longer, which provides indirect evidence for the prompt updates. In this subsection, we present additional evidence by monitoring the percentage change in various linkage networks over time. This analysis allows us to assess the dynamics and updates of these networks, providing further support for the timeliness and prompt updates of the news co-mention linkage.

We calculate the change rate of a given network matrix at week t, denoted as $%Change_t$, to measure the speed of the linkage change in the week. Speci cally:

where \mathbf{M}_{t} is the network matrix at week t, and $\mathbf{M}_{t(i,j)}$ is the element at row i and column j of the network matrix.

We then compute the weekly change rate time series for each linkage type. The summary statistics of the change rate of each linkage are reported in Table 8. The results show that the news co-mention matrix updates much faster than other types of linkages. Speci cally, on average, there is a 16.07% change in co-mention linkages from one week to the next. This proportion is higher than the change rate of analyst linkages (3.61%), statistical linkages (10.53%), and concept linkages (1.27%), which are also updated on a weekly basis. Furthermore, the change rates of industry, geographic, and technology linkages are much lower since their network matrices update on a yearly basis. The mean ratios for these linkages are almost 0, indicating very little change from one week to the next.

These results provide direct evidence that the news co-mention linkage updates more quickly compared to other economic linkages. This faster update speed suggests that the co-mention linkage method is more e ective at capturing and identifying changes in linkages among rms in a timely manner.

	Mean	Std	Min	Median	Max
News co-mention	16.07%	0.20	0.94%	13.61%	316.70%
Analyst	3.61%	0.03	0.11%	2.76%	42.92%
Industry	0.12%	0.01	0.00%	0.00%	21.85%
Geographic	0.10%	0.01	0.00%	0.00%	7.43%
Technology	0.64%	0.05	0.00%	0.00%	42.09%
Statistical	10.53%	0.12	1.69%	8.71%	110.17%
Concept	1.27%	0.08	0.00%	0.48%	116.23%

Table 8: Summary statistics of the linkage change rate

This table reports the summary statistics of the change rate of six economic linkages, including the news co-mention linkage, shared-analyst linkage, industry linkage, geographic linkage, technology linkage, statistical linkage, and concept linkage. The news co-mention linkage is based on *same_sentence* type under the 3-month identic cation window. The industry linkage is based on the Shenwan-1 classi cation. The geographic linkage is at the province level. The change rate of a certain linkage type in one week is computed as below: subtract the network matrix last week from the matrix this week, then take the absolute value of all the elements of the di erence matrix and make a summation, divided by the sum of all elements of the network matrix last week. The sample period of news co-mention linkage is 2012-2020. The sample period of concept linkage is 2006-2020.

5 Mechanism

In this section, we further investigate the mechanism behind this news-based momentum spillover. There are two main theories that attempt to explain the cross- rm momentum anomaly. The rst one is the investors' limited attention theory, which has gained widespread acceptance as an explanation for various types of cross-rm predictability.²² According to the theory, gathering information about peer rms requires additional e ort and attention from investors. As a result, investors may overlook or underestimate the importance of such information, leading to a lead-lag e ect between the stock returns of peer rms and the focal rm. In recent years, a new behavioral-based psychological barrier theory has been proposed by Huang et al. (2021) as an alternative explanation of the momentum spillover e ect. This theory argues that, due to the cognitive bias of anchoring, investors tend to react slowly to positive news from peer rms when the focal stock price is close to its 52-week high. This is because they believe that the stock price has already reached its peak and is unlikely to increase further. Similarly, investors also tend to respond slowly to negative news from peer rms when the focal stock price is far from its 52-week high. ²³This behavior creates a lead-lag e ect between the returns of peer rms and the focal rm. We will examine the two explanations respectively.

5.1 Limitied Attention

Since it is di cult to quantify investors' attention precisely, the literature often relies on indirect metrics to evaluate limited attention. In particular, if the lead-lag e ect in stock returns is a result of investors' inattention to peer rms' information, we would expect to observe a stronger predictive power of the news co-mention

²²In fact, most previous studies attribute the momentum spillover e ect to investors' limited attention to news from peer rms. See Burt and Hrdlicka (2021) for a detailed list of these studies.

²³In a di erent context, Hung et al. (2022) linked the aggregate 52-week high to limited attention. They argued that a higher market level 52-week high is associated with a large amount of good rm-speci c information in the market, which, in turn, may

momentum when it is more challenging for investors to access information about peer rms.

There are several methods to measure investors' attention. Typically, rms with greater analyst coverage, larger market capitalization, and higher institutional holdings are expected to attract more attention from investors (Du et al., 2022), and information about these rms is more readily available and accessible to investors. In this case, investors are less likely to overlook important information. As a result, the stock price of the focal rm will respond more promptly to relevant information, thereby weakening the lead-lag e ects between peer rms. Moreover, in the case of less opaque rms, the speed of information di usion tends to be faster. This means that investors can obtain information about these rms more easily, which in turn reduces the cross- rm momentum e ect among less opaque rms. Furthermore, the complexity of the network can also serve as a measure to assess investors' attention. According to research by Zhu (2019), the predictability of cross- rm momentum diminishes among rms with more intricate linkage networks. This is because it becomes more challenging for investors to understand and identify the peers of rms within a complicated network. The fact that a complex environment may lead agents to fail to account for the informational content is also discussed in theoretical studies, such as Mondria et al. (2022).

To obtain a more robust conclusion, we utilize all the aforementioned indirect proxies to measure investor attention. We employ the total number of analysts covering the rm (# Analysts), and the total number of analyst reports about the rm (# Reports) in the year to capture analyst attention. Additionally, we use the log oat value of the rm to measure its size. The institutional holding ratio (%Institution) is computed as the proportion of shares held by institutional investors to the total shares of the listed rm. The opacity indicator of one rm, denoted as *OPACITY*, is computed as the past-three-year sum of the absolute value of annual discretionary accruals (DISACC) (Hutton et al., 2009). For the network complexity, we follow Ali and Hirshleifer (2020) and use the degree centrality (i.e., the number of peer rms, denoted as # Peers) as the proxy variable.

At the end of each trading week, we create dummy variables for each of the six proxy variables. Each of the dummy variables takes a value of one if the corresponding variable value for a focal rm is higher than the median value of the sample, and zero otherwise. For example, the dummy variable based on the number of analysts *Dummy#Analysts* equals one if the number of analysts of one rm is higher than the sample median, and 0 otherwise. For each of the dummy variables, we add an interaction term between it and the news co-mention peer rm return *Sentence_3_Rtn* to the Fama-MacBeth regression speci ed in subsection 4.2.

Table 9 presents the results of the Fama-MacBeth regressions with the interaction terms. Columns 1-4 show that the coe cient of the interaction terms $Dummy#Analysts \times Sentence_3_Rtn$, $Dummy#Reports \times Sentence_3_Rtn$, $DummyValues \times Sentence_3_Rtn$, and $Dummy%Institution \times Sentence_3_Rtn$ are all signi cantly negative. This suggests that the predictive power of the news co-mention momentum is weaker among rms with higher analyst coverage, more analyst reports, larger oat values, and higher institutional holder proportions. From columns 5-6, the interaction terms $DummyOPACITY \times Sentence_3_Rtn$ and $Dummy#Peers \times Sentence_3_Rtn$ are both positive and signi cant. This suggests that the news co-mention momentum is stronger among rms that are more opaque (whose information is less transparent) and rms with a higher network degree.

	1	2	3	4	5	6
Sentence_3_Rtn	0.040	0.042	0.054	0.041	0.004	0.022
	(6.19)	(6.05)	(6.17)	(6.14)	(1.11)	(5.04)
Dummy#Analysts × Sentence_3_Rtn	-0.014					
	(-2.13)					
Dummy#Reports × Sentence_3_Rtn		-0.024				
		(-3.17)				
DummyValues × Sentence_3_Rtn			-0.052			
			(-4.97)			
Dummy%Institution × Sentence_3_Rtn				-0.013		
				(-2.25)		
DummyOPACITY × Sentence_3_Rtn					0.041	
					(4.75)	
Dummy#Peers × Sentence_3_Rtn						0.033
						(3.87)
Intercept	-0.032	-0.032	0.007	-0.008	-0.007	-0.009
	(-9.69)	(-9.35)	(1.94)	(-2.99)	(-1.91)	(-4.04)
Control	YES	YES	YES	YES	YES	YES
Avg. R Square	0.096	0.097	0.094	0.093	0.093	0.094
Total Obs.	507665	507665	507665	507665	507665	507665
Avg. Obs.	1094	1094	1094	1094	1094	1094

Table 9: Limited attention and FM regressions

This table reports the results of Fama-MacBeth regressions, where interaction terms between the news co-mention peer rm return and di erent limited attention dummy proxies are included. The news co-mention momentum is based on *same_sentence* type under the 3-month identi cation window. The dependent variable is the stock excess return in the next week. The independent variables include the interaction term between the news co-mention peer rm return *Sentence_3_Rtn* and the six dummy variables. The six proxy variables for limited attention are the number of analysts (#*Analysts*), the number of reports (#*Reports*), the oat value (*Values*), the proportion of institutional holders (%*Institution*), the opacity indicator (*OPACITY*) (Hutton et al., 2009), and the network complexity (#*Peers*). We then de ne corresponding dummy variables for these six proxy variables at the end of each trading week. For each dummy, we let it be one if the corresponding variable value of a focal rm is higher than the sample median, and zero otherwise. For example, *Dummy# Analysts* equals one if the number of analysts of one rm is higher than the cross-sectional sample median, else zero. The sample period is 2012-2020. The control variables include the rm size (taking logarithms), book-to-market ratio, and stock return in the past week. All independent variables are shown in parentheses.

We also perform a portfolio analysis by dividing all sample stocks into two groups (high group for dummy=1 and low group for dummy=0) based on each of the six dummy variables created earlier. Subsequently, we sort all stocks within each of the two groups into quintiles based on their values of $Sentence_3_Rtn$. For each of the two groups, we construct a long-short portfolio by purchasing stocks from the highest quintile and selling stocks from the lowest quintile.

Table 10 reports the portfolio grouping sorting results. From columns 1-4, the long-short portfolio exhibits

a higher mean excess return and alpha for the group with fewer analysts, fewer reports, smaller size, and lower institutional holder proportions. These di erences are statistically signi cant, except for the institutional holding one. From columns 5-6, the long-short portfolio exhibits a higher mean excess return and alpha for the group with a higher degree of opacity and a higher network degree.

Overall, both the Fama-MacBeth regressions and the portfolio analysis support the limited attention theory.

Panel A: Exces	ss returns					
	# Analysts	# Reports	Float values	% Institutional	OPACITY	# Peer rms
Higher group	1.37	1.25	0.70	1.64	3.16	2.56
	(4.54)	(4.21)	(2.49)	(5.15)	(5.34)	(4.10)
Lower group	2.44	2.60	3.00	2.06	0.57	1.29
	(5.33)	(5.51)	(5.58)	(4.82)	(2.50)	(5.06)
Higher-Lower	-0.95	-1.14	-2.86	-0.48	2.42	0.94
	(-2.85)	(-3.22)	(-4.34)	(-1.34)	(4.61)	(2.07)
Panel B: CH-4	alphas					
	# Analysts	# Reports	Float values	% Institutional	OPACITY	# Peer rms
Higher group	1.31	1.23	0.70	1.53	3.01	2.41
	(4.20)	(4.05)	(2.32)	(4.84)	(5.55)	(4.04)
Lower group	2.36	2.50	2.82	2.02	0.58	1.28
	(5.41)	(5.50)	(5.63)	(4.74)	(2.36)	(5.02)
Higher-Lower	-0.90	-1.04	-2.51	-0.47	2.27	0.89
	(-2.96)	(-3.26)	(-4.27)	(-1.25)	(4.70)	(1.92)

Table 10: Limited attention and portfolio sorting

This table reports the portfolio sorting results for di erent attention groups. The news co-mention momentum is based on *same_sentence* type under the 3-month identi cation window. At the end of each trading week, we group all sample stock into two groups according to the six proxy variables for limited attention. The six proxy variables for limited attention are the number of analysts (#*Analysts*), the number of reports (#*Reports*), the oat value (*Values*), the proportion of institutional holders (%*Institution*), the opacity indicator (*OPACITY*) (Hutton et al., 2009), and the network complexity (#*Peers*). For each proxy variable, the group with variables larger than the median value is the higher group, while the group with variables smaller than the median is the lower group. We then conduct the portfolio sorting analysis according to *Sentence_3_Rtn* for the two groups respectively. The long-short portfolio of the higher group involves buying the highest quintile and selling the lowest quintile. The sample period is 2012-2020. Panel A shows the excess returns of the long-short portfolios, while Panel B reports the CH-4 adjusted alphas. All returns alphas are converted to monthly using compound interest. Newey and West (1987) adjusted t-statistics are shown in parentheses. Returns/alphas of Higher-Lower portfolios with t-statistics higher/lower than 2.00/-2.00 are highlighted in bold.

5.2 Psychological Barrier

To examine the psychological barrier theory, we adopt the method used in Huang et al. (2021) and de ne the nearness to the 52-week high as the ratio of the closing price at the end of the trading week to the maximum daily closing price observed over the past 12 months.

Panel A: Excess m	nean returns				
	CR1	CR2	CR3	CR4	CR5
PRC1	0.43	1.11	0.99	1.20	0.53
	(0.40)	(1.08)	(1.03)	(1.23)	(0.55)
PRC2	0.55	0.83	1.09	0.49	0.92
	(0.55)	(0.94)	(1.17)	(0.55)	(1.07)
PRC3	0.39	0.74	0.74	0.97	0.49
	(0.46)	(0.81)	(0.80)	(1.12)	(0.59)
PRC4	0.24	1.36	1.03	0.39	0.38
	(0.24)	(1.50)	(1.18)	(0.44)	(0.47)
PRC5	0.72	1.04	0.82	-0.08	1.91
	(0.61)	(0.93)	(0.80)	(-0.09)	(1.73)
PRC5- PRC1	0.25	-0.07	-0.16	-1.27	1.38
	(0.29)	(-0.07)	(-0.20)	(-1.68)	(1.39)
Dr	art CC Dort 11			1.48	
PL	011.55-P011.11			(1.57)	
Panel B: CH-4 adj	usted alphas				
	CR1	CR2	CR3	CR4	CR5
PRC1	-0.08	0.57	0.39	0.80	0.00
	(-0.08)	(0.54)	(0.41)	(0.79)	(0.00)
PRC2	-0.05	0.17	0.68	0.03	0.38
	(-0.05)	(0.18)	(0.75)	(0.04)	(0.44)
PRC3	0.23	0.30	0.47	0.75	0.31
	(0.27)	(0.34)	(0.51)	(0.87)	(0.39)
PRC4	0.10	1.05	0.57	0.38	0.16
	(0.11)	(1.16)	(0.69)	(0.43)	(0.19)
PRC5	1.01	1.06	0.64	-0.13	1.75
	(0.91)	(0.98)	(0.62)	(-0.14)	(1.70)
PRC5- PRC1	1.05	0.49	0.25	-0.92	1.74
	(1.21)	(0.48)	(0.32)	(-1.17)	(1.98)
D	art EE Dort 11			1.83	
PC	011.00-2011.11			(2.04)	

Table 11: Psychologial barrier and news co-mention momentum

This table reports the double sorting performance according to the nearness to the 52-week high (*PRC*) and the news co-mention peer rm return (denoted as *CR* in the table). The news co-mention momentum is based on *same_sentence* type under the 3-month identi cation window. Following Huang et al. (2021), at the end of each trading week, all sample stocks are independently sorted into 5 5 portfolios based on *CR* and *PRC*. The portfolios are held for one week and rebalanced weekly. Stocks are equal-weighted within each group. We also present the di erences in returns between the corner portfolios. For example, Port.55 denotes the portfolio within the CR5 and PRC5 group, and Port.11 denotes the portfolio within the CR1 and PRC1 group. The sample period is 2012-2020. Panel A shows the excess returns of the long-short portfolios, while Panel B reports the CH-4 adjusted alphas. All returns alphas are converted to monthly through compound interest. Newey and West (1987) adjusted t-statistics are shown in parentheses. Returns/alphas of long-short portfolios with t-statistics higher than 2.00 are highlighted in bold.

Then, we construct 5×5 double-sorting equal-weighted portfolios independently based on the nearness to the 52-week high (denoted as *PRC* later) and the news co-mention peer rm return (*Sentence_3_Rtn*, denoted as *CR* in Table 11). If the psychological barrier theory holds, we would expect to observe increasing portfolio returns as we move up the ranks of *PRC*, and the long-short portfolio should exhibit highly signi cant positive returns.

Table 11 reports the mean excess returns and CH-4 adjusted alphas of the double-sorting portfolios. In

contrast to the ndings of Huang et al. (2021), who examined the US stock market, we do not observe an increasing trend in portfolio alpha with the ranks of *PRC* within each *CR* quintile. This result indicates that the psychological barrier theory cannot explain the momentum spillover e ect in the Chinese stock market. The reason could be that the 52-week high e ect itself is not empirically supported in the Chinese stock market. As documented in Hou et al. (2023), the 52-week high e ect in China only leads to a modest adjusted alpha of 0.32%, which is not statistically signi cant. Moreover, the study conducted by Zhang et al. (2019) in the Chinese stock market found evidence of a 52-week low e ect (i.e., stocks with prices near their 52-week lows tend to exhibit higher future returns), which is the opposite of the traditional 52-week high e ect documented by George and Hwang (2004).

6 Futher Analysis and Robustness Checks

6.1 Longer Inverstment Horizons

In the previous portfolio sorting analysis, we hold the portfolio for one week and perform weekly rebalancing. In this subsection, we extend the analysis to examine the predictive power of the news co-mention momentum over longer investment horizons, including holding periods of 2, 3, 4, 8, 12, 24, and 36 weeks. To address the issue of inconsistent rebalancing frequency (which remains weekly) and holding period, we adopt the overlapping portfolio method, as described in Jegadeesh and Titman (1993) and Eisdorfer et al. (2022). We calculate the equal-weighted average of these overlapping portfolios. For instance, with a 2-week holding period, at the end of each trading week, we allocate only half of the total position to construct the portfolio based on the strategy for that week. The portfolio is then held for 2 weeks. In the following week, we use the other half of the total position to construct the portfolio for the next week.

Table 12 shows the performance of long-short portfolios with longer holding periods for di erent momentum spillover e ects. Generally, the predictive power of each cross- rm momentum decreases as the holding period increases. For the news co-mention momentum, as the holding period increases from 1-week to 2-week and then to 3-week, the long-short mean excess return decreases from 1.94% (t-statistic=5.33) to 1.56% (t-statistic=4.95) and 1.18% (t-statistic=4.36) respectively. The noticeable decrease in the strength of the news co-mention momentum under longer investment horizons can be attributed to the fast update of the news co-mention linkage. Peer rms identi ed through news co-mention are time-sensitive, and their relationships could be valid only in the short term. Therefore, investment strategies based on news co-mention momentum should be executed within a short-term timeframe. However, linkages such as the shared-analyst, industry, and concept linkages typically remain relevant and e ective over the longer term, leading to a more gradual decrease in the strength of the corresponding momentum spillover e ects. Despite this, the news co-mention momentum strategy continues to outperform other momentum spillover e ects in most investment horizons, except for the 4-week and 8-week holding periods. However, it should be noted that the dominance of the news co-mention momentum diminishes as the investment horizon lengthens, which is due to the fast update nature of the news co-mention momentum linkage.

Panel A: Excess ret	urns of long	-short portfo	olios				
	2-week	3-week	4-week	8-week	12-week	24-week	36-week
News co-mention	1.55	1.18	0.87	0.42	0.27	0.19	0.18
	(5.24)	(4.74)	(4.78)	(4.20)	(3.20)	(3.29)	(3.63)
Analyst	1.03	0.86	0.80	0.43	0.26	0.16	0.18
	(5.63)	(5.47)	(5.93)	(4.42)	(3.22)	(2.58)	(3.13)
Industry	0.73	0.62	0.54	0.34	0.22	0.13	0.11
	(4.16)	(4.40)	(4.28)	(3.59)	(2.69)	(2.23)	(1.88)
Geographic	0.41	0.30	0.26	0.12	0.08	0.05	0.07
	(5.05)	(4.43)	(4.31)	(2.96)	(2.05)	(1.80)	(2.59)
Technology	0.50	0.34	0.27	0.20	0.14	0.13	0.12
	(3.08)	(2.47)	(2.30)	(2.22)	(1.84)	(2.37)	(2.30)
Statistical	1.32	1.10	0.90	0.51	0.25	0.03	0.14
	(5.22)	(4.83)	(4.48)	(3.24)	(1.91)	(0.27)	(1.79)
Concept	0.65	0.68	0.54	0.32	0.19	0.17	0.19
	(2.56)	(3.36)	(2.66)	(2.30)	(1.73)	(2.22)	(3.13)
Panel B: CH-4 adju	isted alphas	long-short p	ortfolios				
	2-week	3-week	4-week	8-week	12-week	24-week	36-week
News co-mention	1.56	1.18	0.90	0.42	0.27	0.18	0.19
	(4.95)	(4.36)	(4.36)	(4.19)	(3.21)	(3.35)	(3.69)
Analyst	1.05	0.86	0.81	0.42	0.22	0.10	0.14
	(5.29)	(5.09)	(5.76)	(4.02)	(2.65)	(1.73)	(2.47)
Industry	0.76	0.65	0.58	0.32	0.19	0.08	0.07
	(4.00)	(4.26)	(4.36)	(3.20)	(2.33)	(1.42)	(1.26)
Geographic	0.42	0.30	0.27	0.12	0.07	0.04	0.05
	(4.95)	(4.27)	(4.29)	(2.74)	(1.72)	(1.35)	(1.93)
Technology	0.53	0.34	0.28	0.17	0.12	0.09	0.08
	(3.05)	(2.36)	(2.25)	(1.89)	(1.49)	(1.66)	(1.61)
Statistical	1.36	1.14	0.97	0.53	0.27	0.00	0.10
	(4.73)	(4.39)	(4.37)	(3.09)	(2.06)	(0.00)	(1.36)
Concept	0.58	0.66	0.55	0.34	0.19	0.15	0.17
	(2.20)	(2.98)	(2.48)	(2.30)	(1.70)	(1.87)	(2.96)

Table 12: Momentum performances over longer investment horizons

This table reports the long-short returns and alphas of the news co-mention momentum, analyst momentum, industry momentum, geographic momentum, technology momentum, statistical momentum, and concept momentum over longer investment horizons. The news co-mention momentum is based on *same_sentence* type under the 3-month identi cation window. The industry momentum is based on the Shenwan-1 classi cation. The geographic momentum is at the province level. We convert the holding period from 1-week to 2-, 3-, 4-, 8-, 12-, 24-, and 36-week. We follow the method in Jegadeesh and Titman (1993) and Eisdorfer et al. (2022) to have overlapping portfolios. The sample period of news co-mention momentum is 2012-2020. The sample period of concept momentum is Aug. 2016-2020. The sample periods for other momentums are 2006-2020. We report the equal-weighted average of these overlapping portfolios. Panel A shows the excess returns of portfolios, while Panel B shows the intercepts of the regression of the returns on CH-4 factors (Liu et al., 2019) (market, size, value, abnormal turnover rate). All returns and alphas are converted into monthly. Newey and West (1987) adjusted t-statistics are shown in parentheses.

6.2 Heterogeneity of Momentum Spillover E ects for SOEs and Non-SOEs

In this subsection, we examine the heterogeneity of momentum spillover e ects based on ownership di erences. State-owned enterprises (SOEs) play a critical role in shaping the market dynamics and contributing to the overall economy. SOEs di er from non-state-owned enterprises (non-SOEs) due to their considerations of not only corporate economic performance but also political objectives (Jiang and Kim, 2020; Leippold et al., 2022).

These additional factors often in uence the stock prices of SOEs. We conduct separate portfolio analyses for SOEs and non-SOEs.

At the end of each trading week, we divide the sample stocks into two groups: state-owned enterprises (SOEs) and non-state-owned enterprises (non-SOEs). For each linkage type, we then perform separate portfolio sorting analyses for SOEs and non-SOEs.

Panel A: Excess r	eturns of long-short	portfolios					
	News co-mention	Analyst	Industry	Geographic	Technology	Statistical	Concept
non-SOE	2.48	0.82	0.59	0.34	0.28	1.29	0.60
	(5.90)	(3.45)	(2.56)	(2.82)	(1.17)	(4.29)	(1.68)
SOE	0.73	0.81	0.83	0.30	0.66	1.20	0.94
	(2.81)	(3.01)	(3.57)	(2.03)	(3.10)	(3.36)	(2.87)
non-SOE - SOE	1.73	0.01	-0.24	0.04	-0.38	0.09	-0.34
	(5.23)	(0.05)	(-1.31)	(0.27)	(-1.56)	(0.39)	(-1.17)
Panel B: CH-4 ac	ljusted alphas of lor	ig-short po	rtfolios				
	News co-mention	Analyst	Industry	Geographic	Technology	Statistical	Concept
non-SOE	2.38	0.84	0.65	0.36	0.31	1.23	0.50
	(5.69)	(3.24)	(2.67)	(3.08)	(1.27)	(3.51)	(1.32)
SOE	0.71	0.77	0.75	0.29	0.65	1.17	0.93
	(2.48)	(2.53)	(3.02)	(1.72)	(2.92)	(2.84)	(2.63)
non-SOE - SOE	1.66	0.07	-0.11	0.07	-0.33	0.06	-0.43
	(4.67)	(0.35)	(-0.58)	(0.45)	(-1.37)	(0.26)	(-1.44)

Table 13: Heterogeneity of momentum spillover e ects for SOEs and Non-SOEs

This table reports the di erences in momentum spillover e ects between SOEs and non-SOEs. The news co-mention momentum is based on *same_sentence* type under the 3-month identi cation window. The industry momentum is based on the Shenwan-1 classi cation. The geographic momentum is at the province level. At the end of each trading week, we divide the sample stocks into SOEs and non-SOEs. Then within each sample group, we sort all stocks into quintiles according to the peer rm returns based on each linkage type. The long-short portfolio is buying the highest quintile and selling the lowest quintile. Non-SOE - SOE indicates the taking the di erence of the time series long-short returns and alphas between the non-SOE group and SOE group. Within each quintile group, the stocks are equally weighted. All portfolios are held for one week and are rebalanced weekly. The sample period of news co-mention momentum is 2012-2020. The sample period of concept momentum is Aug. 2016-2020. The sample periods for other momentums are 2006-2020. Panel A shows the excess returns of portfolios, while Panel B shows the intercepts of the regression of the returns on CH-4 factors (market, size, value, abnormal turnover rate). All returns and alphas are converted into monthly. Newey and West (1987) adjusted t-statistics are shown in parentheses. Spreads between non-SOEs and SOEs with t-statistics greater than 2.00 are highlighted in bold.

Table 13 reports the di erences in momentum spillover e ects between SOEs and non-SOEs. Interestingly, among the various momentum spillover e ects examined, only the news co-mention momentum demondi erences in returns due to the ownership between non-SOEs and SOEs are not statistically signi cant for other momentum spillovers. It is possible that the observed results are due to media bias. Given that SOEs in China are politically related and often have political objectives (Jiang and Kim (2020) and Schweizer et al. (2020)), it is plausible that the news co-mentioning about SOEs may be less precise in capturing peer rm relationships compared to non-SOEs.²⁴

6.3 Robustness Checks

In this subsection, we conduct several robustness tests to examine the robustness of our ndings. Speci cally, we consider the impact of transaction costs and the shell e ect on our results. In addition, we conduct a series of robustness checks for the unifying e ect of the news co-mention momentum.

6.3.1 Transaction Cost

The portfolio sorting analysis is conducted on a weekly basis in this study. Compared to monthly rebalancing, the weekly approach involves more frequent trading activities, making it important to consider transaction costs in real-world investment. To address this concern, follow Fan et al. (2021) to set the transaction cost at 16 bps per trade (buying and selling combined).²⁵

Table 15 in Appendix C presents the portfolio sorting results of each type of cross- rm momentum, taking into account the 16 bps transaction cost. The trading strategy based on the news co-mention momentum remains signi cantly pro table after taking the transaction costs into account. The long-short portfolio generates an average excess return of 1.29% (t-statistic=3.56) and a CH-4 adjusted alpha of 1.21% (t-statistic=3.42) per month.

However, other momentum spillover e ects do not generate pro table trading strategies after taking the transaction costs into account. The analyst momentum strategy yields a long-short mean excess return of 0.21% and an alpha of 0.20%, both of which are statistically insigni cant. Similarly, the industry momentum e ect does not generate pro table trading strategies once we consider the transaction cost. The statistical momentum strategy, which initially showed strong performance without considering transaction costs, experiences a signi cant decline both economically and statistically when transaction costs are taken into account. The magnitude of the excess mean return decreases considerably, and the CH-4 adjusted alpha becomes statistically insigni cant at 0.69% with a t-statistic of 1.74. Furthermore, when considering the 16 bps transaction cost, the trading strategies based on geographic momentum, technology momentum, and concept momentum exhibit negative long-short returns and CH-4 adjusted alphas.

²⁴In China, SOEs often have closer ties to the government and are subject to more government oversight compared to non-stateowned enterprises (non-SOEs). The government may strategically use media outlets for economic and political purposes, leading to potential biases in news coverage.

²⁵In the context of stock trading in China, transaction costs typically consist of three components. Firstly, there is a stamp duty, which is levied on the total transaction amount at a rate of 10 bps. It is important to note that the stamp duty is only imposed on sellers. The second component is the transfer fee, which amounts to 1 bps for both buying and selling transactions in the Shanghai Stock Exchange. This fee is applicable to stocks with a price of 20 CNY per share. Lastly, there is the trading commission, which is the fee paid by investors to brokers for executing their trades. The trading commission is subject to a maximum limit of 3 bps of the transaction amount. However, the typical rates for trading commission rates compared to individual investors. Setting the transaction cost at 16 bps implicitly assumes a conservative approach by considering a turnover ratio of 100% for each rebalancing period.

6.3.2 Dropping Shell Firms

According to Liu et al. (2019), back-door listings through reverse mergers are common in China due to the strict and costly IPO process. This can lead to the presence of shell rms with in ated market values and biased stock returns that are disconnected from their underlying fundamentals. To address this potential issue and ensure the robustness of our analysis, we exclude shell rms from the dataset and re-evaluate the portfolio sorting analysis for each cross- rm momentum.

At the end of each trading week, we sort all sample stocks except for the ST shares and stocks with the bottom 30% capitalization based on their peer rm returns. The stocks are then divided into quintiles, and within each quintile, they are equally weighted. The long-short portfolio strategy involves buying the stocks in the highest quintile and selling the stocks in the lowest quintile. These portfolios are held for one week and rebalanced weekly.

Table 16 in Appendix C reports the excess reuturns and CH-4 adjusted alphas after droppping the shell rms. Our main results are robust to the exclusion of shell rms. Despite a small decrease in the predictive power, the news co-mention momentum continues to dominate other momentum spillover e ects. To be speci c, the news co-mention momentum strategy generates a statistically signi cant long-short average return and CH-4 adjusted alpha of 1.00% (t-statistic=3.87) and 0.95% (t-statistic=3.63), respectively. The analyst, industry, geographic, and statistical momentum also exhibit a slight decrease in their predictive power but remain statistically signi cant. However, the technology and concept momentum lose their predictability after excluding shell rms from the analysis.

6.3.3 Robustness of the Unifying Effect

In this section, we perform a series of robustness checks to further validate the unifying e ect of the news co-mention momentum.

Firstly, we conduct the spanning tests without CH-4 factors. In this way, we can exclude the in uence of the CH-4 factors and better examine the explanatory power of each momentums spillover factor. Table 17 in Appendix C shows that the news co-mention factor alone is su cient to explain all other momentum spillover e ects. However, none of the other six factors demonstrate the same level of explanatory power as the news co-mention factor across the four identi cation windows.

In the main body of the paper, we primarily focus on the industry linkage based on the Shenwan-1 classi - cation. Next, we test the unifying e ect of the news co-mention momentum over the industry momentum based on alternative industry classi cations, including the three levels of Shenwan classi cation (Shenwan-1, -2, -3), two levels of CSRC classi cation, (CSRC-1, -2), and three levels of CITIC classi cation (CITIC-1, -2, -3). For brevity, for the news co-mention factors, we only use the Sentence_3 factor. Table 18 in Appendix C reports the alphas of the news co-mention momentum factor from the CH-4+Industry model and the alphas of the eight industry factors from the CH-4+Sentence_3 model. The news co-mention momentum factor always has a signi cantly positive alpha given all the candidate industry factors. On the contrary, all the eight industry momentum factors are explained by the news co-mention momentum factor, and their alphas are negative when the Sentence_3 factor is included.

The full sample period analyzed is from 2012 to 2020 in the main body of the paper. This is primarily due to the limited availability of news data in the earlier years. In contrast, most of the other momentum factors were

analyzed from 2006 to 2020. As a robustness check, we consider an alternative sample period from 2006-2020,²⁶ and the results are reported in Table 19 from Appendix C. The news co-mention momentum factor cannot be explained by any of the other factors, and it always has a signi cantly positive alpha in all speci cations. In contrast, all the other factors can be explained by the news co-mention momentum factor, as their alphas are either insigni cant or even exhibit a negative sign when the Sentence_3 factor is included.

Finally, we turn to *same_article* link identi cation strategy when conducting spanning tests. Speci cally, we construct the news co-mention momentum factor based on *same_article* type under identi cation windows of 3, 6, 9, and 12 months. The spanning test results are shown in Table 20 from Appendix C. Our main results are robust to the choice of news linkage identi cation strategy.

7 Conclusions

In recent years, there has been a surging interest in the extraction and utilization of soft information contained within news articles. This paper utilizes a large dataset of millions of Chinese news articles and applies both article co-mentioning and sentence co-mentioning strategies to identify economic linkages among rms. Our work demonstrates the bene ts of utilizing big alternative big data by showing that economic linkages identi ed from business news contain rich and valuable information. We nd that the news co-mention momentum spillover uni es all di erent forms of momentum spillover e ects that have been previously studied in the Chinese stock market. These include shared-analyst momentum, industry momentum, geographic momentum, and concept momentum. This result remains robust even when considering di erent choices of news-linkage identi cation strategy, alternative de nitions for competitor links, and various sample periods.

To further explore the source of this unifying e ect, we examine the di erences in the information content of news-implied links and other linkages. We nd that the news-implied linkages are more comprehensive compared to other types of linkage proxies. Speci cally, news co-mentioning incorporates both industry information and valuable non-industry linkage information, which contributes to its strong predictive power for future returns. Furthermore, news exhibits the advantage of prompt updates, allowing us to timely identify changes in linkages among rms.

We also investigate the mechanism behind the news-based momentum spillover. We nd that the news co-mention momentum is weaker among rms with higher analyst coverage, more analyst reports, larger oat values, higher institutional holder proportions, less opacity, and less linkage complexity, supporting the limited attention explanation.

In the last, we wish to point out that the trading strategy based on news co-mention momentum spillover e ect is pro table when considering real-world transaction costs. However, due to the quick update nature of news, investment strategies based on news co-mention momentum should be executed within a short-term timeframe.

²⁶The news database starts from 2006.

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Appendix

Appendix A. Descriptive Statistics for the News Data

Year	Mean	Std.	Max	Min
2006	97.23	59.60	252	1
2007	88.04	61.19	337	1
2008	70.72	28.88	152	3
2009	57.83	27.19	137	12
2010	61.06	33.66	135	1
2011	72.23	41.87	194	1
2012	84.29	52.16	183	1
2013	243.08	185.71	657	1
2014	336.04	212.59	699	1
2015	236.23	158.75	589	1
2016	197.29	131.52	493	1
2017	126.35	89.32	359	1
2018	156.37	105.31	467	1
2019	136.68	94.55	464	1
2020	334.87	313.69	1685	1

Table 14: Descriptive statistics for the news data

This table summarizes the daily number of news items for each year from 2006 to 2020.

Appendix B. An Eexample of Analysts' Research Directions

Another name for analysts in China is industry researchers. This name can better re ect the fact that most Chinese analysts focus on one speci c industry and are accordingly grouped in advance. Therefore, the analyst

Appendix C. Tables of Robustness Checks

Panel A: Exc	cess returns						
	News co-mention	Analyst	Industry	Geographic	Technology	Statistical	Concept
1 (Low)	0.09	0.53	0.70	0.87	0.89	0.27	-0.96
	(0.11)	(0.72)	(0.95)	(1.18)	(1.19)	(0.36)	(-1.03)
2	0.42	0.63	0.87	0.99	1.01	0.70	-0.88
	(0.52)	(0.89)	(1.18)	(1.34)	(1.35)	(0.91)	(-0.97)
3	0.46	0.96	1.09	1.13	1.29	1.07	-0.59
	(0.58)	(1.34)	(1.43)	(1.51)	(1.71)	(1.40)	(-0.66)
4	0.74	1.38	1.25	1.19	1.27	1.45	-0.43
	(0.88)	(1.92)	(1.64)	(1.59)	(1.67)	(1.93)	(-0.49)
5 (High)	2.03	1.39	1.47	1.22	1.47	1.63	-0.36
	(2.25)	(1.87)	(1.91)	(1.58)	(1.90)	(2.17)	(-0.42)
5-1	1.29	0.21	0.12	-0.29	-0.06	0.71	-0.04
	(3.56)	(0.83)	(0.53)	(-2.36)	(-0.31)	(2.07)	(-0.11)
SpearmanR	1.00	1.00	1.00	1.00	0.90	1.00	1.00
P value	0.00	0.00	0.00	0.00	0.04	0.00	0.00
Panel B: CH	-4 adjusted alphas						
	News co-mention	Analyst	Industry	Geographic	Technology	Statistical	Concept
1 (Low)	-0.23	0.18	0.33	0.49	0.47	-0.06	-0.89
	(-0.28)	(0.24)	(0.44)	(0.66)	(0.61)	(-0.07)	(-0.92)
2	0.14	0.28	0.48	0.60	0.62	0.31	-0.85
	(0.17)	(0.38)	(0.63)	(0.79)	(0.81)	(0.39)	(-0.90)
3	0.13	0.64	0.68	0.75	0.97	0.64	-0.58
	(0.17)	(0.88)	(0.88)	(0.98)	(1.26)	(0.84)	(-0.63)
4	0.42	1.08	0.84	0.77	0.89	1.01	-0.46
	(0.51)	(1.49)	(1.10)	(1.01)	(1.15)	(1.33)	(-0.49)
5 (High)	1.62	1.03	1.10	0.83	1.07	1.28	-0.39
	(1.91)	(1.36)	(1.39)	(1.07)	(1.33)	(1.67)	(-0.44)
5-1	1.21	0.20	0.12	-0.30	-0.04	0.69	-0.14
	(3.42)	(0.72)	(0.47)	(-2.29)	(-0.18)	(1.74)	(-0.36)
SpearmanR	0.90	0.90	1.00	1.00	0.90	1.00	1.00
P value	0.04	0.04	0.00	0.00	0.04	0.00	0.00

Table 15: Considering transaction costs

This table reports the portfolio sorting performances of the news co-mention momentum, analyst momentum, industry momentum, geographic momentum, technology momentum, statistical momentum, and concept momentum considering a transaction cost of 16 bps (buy and sell combined). The sample stocks include all listed stocks on the main board of the Shanghai Stock Exchange, Shenzhen Stock Exchange, and Growth Enterprise Market (GEM). ST shares are excluded. The news co-mention momentum is based on *same_sentence* type under the 3-month identi cation window. The industry momentum is based on *the Shenwan-1* classi cation. The geographic momentum is at the province level. The sample period of news co-mention momentum is 2012-2020. The sample period of concept momentum is Aug. 2016-2020. The sample periods for other momentums are 2006-2020. At the end of each trading week, all sample stocks are sorted into quintiles based on di erent peer rm returns respectively. Within each quintile group, the stocks are equally weighted. The long-short portfolio involves buying the highest group and selling the lowest group. All portfolios are held for one week and are rebalanced weekly. Panel A presents the excess returns of portfolios, and Panel B shows the intercepts of the regression of the returns on CH-4 factors (Liu et al., 2019) (market, size, value, abnormal turnover rate). All weekly returns and alphas are converted into monthly percentages using compound interest. Spearman R reports the Spearman correlation coe cient between the portfolio return and the serial number for each sorting. Newey and West (1987) adjusted t-statistics are shown in parentheses. Long-short returns/alphas with t-statistics higher than 2.00 are highlighted in bold.

Panel A: Exc	ess returns						
	News co-mention	Analyst	Industry	Geographic	Technology	Statistical	Concept
1	0.43	0.92	0.95	1.10	1.13	0.83	-0.52
	(0.53)	(1.25)	(1.30)	(1.53)	(1.55)	(1.12)	(-0.56)
2	0.78	1.04	1.09	1.19	1.15	1.05	-0.25
	(0.99)	(1.48)	(1.52)	(1.64)	(1.59)	(1.40)	(-0.27)
3	0.84	1.38	1.28	1.29	1.49	1.38	-0.03
	(1.09)	(1.95)	(1.72)	(1.77)	(2.03)	(1.85)	(-0.04)
4	1.04	1.79	1.37	1.39	1.47	1.60	-0.16
	(1.28)	(2.50)	(1.84)	(1.89)	(1.97)	(2.20)	(-0.19)
5	1.43	1.74	1.69	1.44	1.57	1.80	0.06
	(1.73)	(2.34)	(2.22)	(1.90)	(2.08)	(2.42)	(0.07)
5-1	1.00	0.81	0.74	0.33	0.44	0.96	0.58
	(3.87)	(2.96)	(2.73)	(2.34)	(1.84)	(2.63)	(1.52)
SpearmanR	1.00	0.90	1.00	1.00	0.90	1.00	0.90
P value	0.00	0.04	0.00	0.00	0.04	0.00	0.04
Panel B: CH	4 adjusted alphas						
	News co-mention	Analyst	Industry	Geographic	Technology	Statistical	Concept
1	0.17	0.60	0.62	0.79	0.75	0.53	-0.47
	(0.21)	(0.79)	(0.83)	(1.07)	(1.00)	(0.69)	(-0.50)
2	0.52	0.73	0.79	0.86	0.82	0.72	-0.23
	(0.67)	(1.01)	(1.07)	(1.16)	(1.09)	(0.95)	(-0.25)
3	0.56	1.10	0.96	0.97	1.21	1.02	-0.06
	(0.73)	(1.51)	(1.26)	(1.30)	(1.62)	(1.35)	(-0.06)
4	0.76	1.54	0.99	1.04	1.15	1.25	-0.24
	(0.95)	(2.14)	(1.32)	(1.39)	(1.51)	(1.69)	(-0.26)
5	1.12	1.40	1.37	1.10	1.21	1.48	-0.02
	(1.40)	(1.84)	(1.73)	(1.43)	(1.55)	(1.94)	(-0.02)
5-1	0.95	0.80	0.74	0.32	0.46	0.95	0.46
	(3.63)	(2.62)	(2.55)	(2.05)	(1.84)	(2.25)	(1.14)
SpearmanR	1.00	0.90	1.00	1.00	0.90	1.00	0.70
P value	0.00	0.04	0.00	0.00	0.04	0.00	0.19

Table 16: Dropping shell rms

This table reports the portfolio sorting performances of the news co-mention momentum, analyst momentum, industry momentum, geographic momentum, technology momentum, statistical momentum, and concept momentum after dropping all stocks with the bottom 30% capitalization on the main board of the Shanghai Stock Exchange, Shenzhen Stock Exchange, and Growth Enterprise Market (GEM). ST shares are also excluded. The news co-mention momentum is based on *same_sentence* type under the 3-month identi cation window. The industry momentum is based on the Shenwan-1 classi cation. The geographic momentum is at the province level. The sample period of news co-mention momentum is 2012-2020. The sample period of concept momentum is Aug. 2016-2020. The sample periods for other momentums are 2006-2020. At the end of each trading week, all sample stocks are sorted into quintiles based on di erent peer rm returns respectively. Within each quintile group, the stocks are equally weighted. The long-short portfolio involves buying the highest group and selling the lowest group. All portfolios are held for one week and are rebalanced weekly. Panel A is the excess returns of portfolios, and Panel B is the intercepts of the regression of the returns on CH-4 factors (Liu et al., 2019) (market, size, value, abnormal turnover rate). All weekly returns and alphas are converted into monthly percentages using compound interest. SpearmanR reports the Spearman correlation coe cient between the portfolio return and the serial number for each sorting. Newey and West (1987) adjusted t-statistics are shown in parentheses. Long-short returns/alphas with t-statistics higher than 2.00 are highlighted in bold.

Dependent factors Explainoary factors	Sentence_3	Sentence_6	Sentence_9	Sentence_12	Analyst	Industry	Geographic	Technology	Concept
Sentence_3		-0.18	-0.25	-0.34	-0.53	-0.62	0.22	-0.36	-0.54
		(-1.79)	(-1.96)	(-2.45)	(-2.04)	(-2.48)	(1.74)	(-1.54)	(-1.70)
Sentence_6	0.41		-0.09	-0.18	-0.40	-0.51	0.24	-0.30	-0.45
	(3.78)		(-1.26)	(-2.16)	(-1.73)	(-2.29)	(1.97)	(-1.37)	(-1.49)
Sentence_9	0.58	0.17		-0.09	-0.31	-0.44	0.24	-0.25	-0.34
	(4.23)	(2.34)		(-2.01)	(-1.39)	(-2.10)	(1.95)	(-1.20)	(-1.17)
Sentence_12	0.71	0.29	0.13		-0.22	-0.36	0.25	-0.20	-0.27
	(4.66)	(3.39)	(2.72)		(-1.02)	(-1.77)	(2.05)	(-0.96)	(-0.98)
Analyst	1.27	0.86	0.72	0.62		0.11	0.25	0.14	0.07
	(5.07)	(4.30)	(3.80)	(3.43)		(0.82)	(2.01)	(0.90)	(0.32)
Industry	1.43	1.02	0.89	0.79	0.15		0.26	0.12	0.37
	(5.15)	(4.59)	(4.33)	(3.98)	(1.08)		(2.09)	(0.77)	(1.88)
Geographic	1.69	1.32	1.17	1.08	0.60	0.55		0.42	0.46
	(5.19)	(4.61)	(4.17)	(3.90)	(2.42)	(2.40)		(2.05)	(1.38)
Technology	1.59	1.19	1.06	0.97	0.37	0.29	0.28		0.35
	(4.72)	(4.19)	(3.96)	(3.71)	(1.89)	(1.63)	(2.23)		(1.32)
Statistical	1.25	0.83	0.67	0.55	0.20	0.22	0.24	0.22	0.12
	(4.50)	(3.64)	(3.19)	(2.75)	(0.94)	(1.04)	(1.84)	(1.15)	(0.39)
Concept	1.41	0.97	0.77	0.63	0.31	-0.19	0.31	-0.03	
	(4.28)	(3.95)	(3.31)	(3.04)	(1.38)	(-1.02)	(2.03)	(-0.14)	
Non_news	1.23	0.81	0.58	0.45					
	(3.93)	(3.36)	(2.49)	(2.20)					
This table reports the inter- given in subsection 4.1. The	cept (or alpha) of rec column name indice	gressing the time se ates the type of lon	eries of di erent lonç a-short momentum	J-short momentum re return, i.e., the deper	eturns on the M	'S' momentum spi while the index n	illover factor alone. ⁻ ame means the <i>MS</i>	The construction of factor, i.e., the expl	the MS factor is anatory variable.
For example, the alphas of t	he row Sentence_3 in	Judicate the alphas fi	rom the time-series	regressions of long-sh	nort portfolio rei	turns of other mo	mentums on the Sen	tence_3 factor alone	. In the last row,
we include all the six momen	ntum spillover factor	rs that are not relat	ed to news as addit	ional explanatory va	iriables. The nev	vs co-mention mo	mentum factors are	based on same-seni	tence type under

identi cation windows including 3-month, 6-month, 9-month, and 12-month. The industry momentum is based on the Shenwan-1 classi cation, while the geographic momentum is conducted on the province level. We do not consider the customer momentum due to the low data quality. The sample stocks include all listed stocks on the main board of the Shanghai Stock Exchange, Shenzhen

2016-2020. The sample periods for other momentums are 2006-2020. Newey and West (1987) adjusted t-statistics are shown in parentheses. Alphas with t-statistics higher than 2.00 are highlighted Stock Exchange, and Growth Enterprise Market (GEM). ST shares are excluded. The sample period of news co-mention momentum is 2012-2020. The sample period of concept momentum is Aug.

in bold.

Table 17: Spanning tests without CH-4 factors

	1	2	
	Alphas of Sentence_3		CH-4+Sentence_3
CH-4+Shenwan-1	1.35	Alpha of Shenwan-1	-0.53
	(5.08)		(-2.10)
CH-4+Shenwan-2	1.34	Alpha of Shenwan-2	-0.54
	(5.02)		(-2.08)
CH-4+Shenwan-3	1.30	Alpha of Shenwan-3	-0.48
	(5.17)		(-2.11)
CH-4+CSRC-1	1.55	Alpha of CSRC-1	-0.11
	(4.94)		(-0.45)
CH-4+CSRC-2	1.44	Alpha of CSRC-2	-0.64
	(5.38)		(-2.38)
CH-4+CITIC-1	1.48	Alpha of CITIC-1	-0.71
	(5.54)		(-2.66)
CH-4+CITIC-2	1.27	Alpha of CITIC-2	-0.52
	(4.97)		(-2.01)
CH-4+CITIC-3	1.33	Alpha of CITIC-3	-0.59
	(5.48)		(-2.57)

Table 18: Robustness checks based on alternative industry classi cations

This table reports the factor-spanning test results of the industry momentum under di erent classi cation systems. We adopt three levels of Shenwan classi cation (Shenwan-1, -2, -3), two levels of CSRC classi cation, (CSRC-1, -2), and three levels of CITIC classi cation (CITIC-1, -2, -3) to construct the industry momentum factor. The Sentence_3 news co-mention factor is based on *same_sentence* type under the 3-month identi cation window. Column 1 indicates CH-4+ Industry adjusted alphas of the Sentence_3 long-short return based on the eight classi cation methods, respectively, and the row name indicates the speci c type of each CH-4+ Industry model. Column 2 represents the CH-4+ Sentence_3 adjusted alphas of the industry momentum based on the eight classi cation methods. The CH-4 factors (*MKT*, *SMB*, *HML*, *PMO*) are from Liu et al. (2019). The sample stocks include all listed stocks on the main board of the Shanghai Stock Exchange, and Growth Enterprise Market (GEM). ST shares are excluded. The sample period is 2012-2020. Newey and West (1987) adjusted t-statistics are shown in parentheses. Alphas with t-statistics higher than 2.00 are highlighted in bold.

Panel A: Period: 200	6-2020					
	Sentence_3	Analyst	Industry	Geographic	Technology	Statistical
CH-4+Sentence_3		-0.09	-0.06	0.23	0.05	0.44
		(-0.45)	(-0.27)	(1.67)	(0.22)	(1.21)
CH-4+Analyst	0.71		0.11	0.24	0.17	0.70
	(3.61)		(0.76)	(1.83)	(0.99)	(2.18)
CH-4+Industry	0.74	0.15		0.25	0.14	0.78
	(3.54)	(0.96)		(1.92)	(0.87)	(2.25)
CH-4+Geographic	1.01	0.59	0.55		0.44	1.05
	(3.95)	(2.21)	(2.26)		(2.07)	(2.61)
CH-4+Technology	0.85	0.34	0.27	0.27		0.93
	(3.42)	(1.59)	(1.41)	(2.05)		(2.50)
CH-4+Statistical	0.72	0.20	0.23	0.23	0.25	
	(2.96)	(0.90)	(0.97)	(1.69)	(1.24)	
CH-4+Non_news	0.61	0.32	0.22	0.12		
	(2.99)	(1.94)	(1.42)	(0.87)		
Panel B: Period: 201	2-2020					
	Sentence_3	Analyst	Industry	Geographic	Technology	Statistical
CH-4+Sentence_3		-0.44	-0.53	0.25	-0.26	0.17
		(-1.76)	(-2.10)	(1.92)	(-1.11)	(0.43)
CH-4+Analyst	1.18		-0.07	0.31	0.04	0.87
	(4.97)		(-0.43)	(2.55)	(0.22)	(2.54)
CH-4+Industry	1.35	0.34		0.31	0.10	1.06
	(5.08)	(2.01)		(2.59)	(0.60)	(3.01)
CH-4+Geographic	1.62	0.77	0.46		0.40	1.34
	(4.94)	(2.38)	(1.56)		(1.52)	(3.02)
CH-4+Technology	1.49	0.53	0.22	0.33		1.28
	(4.63)	(2.05)	(1.09)	(2.90)		(3.08)
CH-4+Statistical	1.16	0.21	0.00	0.28	0.15	
	(4.17)	(0.81)	(0.01)	(2.24)	(0.63)	
CH-4+Non_news	1.07	0.70	0.52	0.42		

Table 19: Spanning tests of di erent sample periods

This table reports the intercept (or alpha) of regressing the time series of di erent long-short momentum returns on the CH-4 factors (*MKT*, *SMB*, *HML*, *PMO*) (Liu et al., 2019) plus each *MS* momentum spillover factor during sample periods of 2006-2020 and 2012-2020. Panel A reports the results of 2006-2020, while Panel B reports the results of 2012-2020. The concept momentum is excluded from this table since it is only available since 2016. The construction of the *MS* factor is given in subsection 4.1. The column name indicates the type of long-short momentum return (i.e., the dependent variable), while the rows name indicates the augmented CH-4 model (i.e., the original CH-4 model plus a speci c *MS* factor). For example, the alphas in the row CH-4+Sentence_3 are the alphas from the time-series regressions of long-short portfolio returns of other momentum spillover factors (concept momentum is excluded) that are not related to news as additional explanatory variables. For brevity, we only construct the news co-mention momentum factor based on *same_sentence* type under the 3-month identi cation windows here. The industry momentum is based on the Shenwan-1 classi cation, while the geographic momentum is conducted on the province level. We do not consider the customer momentum due to the low data quality. The sample stocks include all listed stocks on the main board of the Shanghai Stock Exchange, Shenzhen Stock Exchange, and Growth Enterprise Market (GEM). ST shares are excluded. Newey and West (1987) adjusted t-statistics are shown in parentheses. Alphas with t-statistics higher than 2.00 are highlighted in bold.

Table 20: Factor spanning tests based on same_article type news co-mention

Panel A: Alphas of factor	r spanning test	S (same_article	type)							
	Article_3	Article_6	Article_9	Article_12	Analyst	Industry	Geographic	Technology	Statistical	Concept
CH-4 + Article_3		-0.05	-0.14	-0.20	-0.44	-0.56	0.24	-0.33	0.18	-0.53
		(-0.62)	(-1.23)	(-1.48)	(-1.82)	(-2.42)	(1.81)	(-1.51)	(0.49)	(-1.77)
CH-4 + Article_6	0.25		-0.09	-0.15	-0.39	-0.53	0.25	-0.32	0.18	-0.51
	(2.72)		(-1.36)	(-1.75)	(-1.74)	(-2.42)	(1.91)	(-1.46)	(0.54)	(-1.79)
CH-4+Article_9	0.41	0.17		-0.07	-0.28	-0.43	0.25	-0.22	0.23	-0.42
	(3.56)	(2.60)		(-1.24)	(-1.33)	(-2.06)	(1.97)	(-1.05)	(0.75)	(-1.49)
CH-4+Article_12	0.51	0.27	0.10	~	-0.21	-0.37	0.26	-0.16	0.27	-0.29
	(3.74)	(3.05)	(1.92)		(-0.99)	(-1.76)	(2.07)	(-0.75)	(0.93)	(-0,-)
CH-4 + Analyst	0.98	0.78	0.67	0.60		0.11	0.24	0.17	0.70	0.06
	(4.85)	(4.63)	(3.83)	(3.46)		(0.76)	(1.83)	(0.99)	(2.18)	(0.28)
CH-4 + Industry	1.13	0.92	0.82	0.76	0.15		0.25	0.14	0.78	0.33
(in the second s	(5.07)	(4.98)	(4.30)	(4.10)	(0.96)		(1.92)	(0.87)	(2.25)	(1.67)
CH-4 + Geographic	1.35	1.18	1.09	1.05	0.59	0.55		0.44	1.05	0.34
-	(4.79)	(4.47)	(3.90)	(3.75)	(2.21)	(2.26)		(2.07)	(2.61)	(0.92)
CH-4 + Technology	1.23	1.04	0.95	0.91	0.34	0.27	0.27	•	0.93	0.25
	(4.59)	(4.38)	(3.77)	(3.62)	(1.59)	(1.41)	(2.05)		(2.50)	(0.98)
CH-4 + Statistical	0.97	0.74	0.60	0.50	0.20	0.23	0.23	0.25		0.04
	(4.04)	(3.53)	(2.94)	(2.65)	(06.0)	(0.67)	(1.69)	(1.24)		(0.14)
CH-4+Concept	1.06	0.83	0.70	0.58	0.29	-0.16	0.35	0.04	0.92	
	(4.07)	(3.94)	(3.45)	(2.91)	(1.30)	(-0.89)	(2.19)	(0.18)	(2.61)	
CH-4+Non news	0 80	0 2 0	051	0 40						
	(3.50)	(3.44)	(2.71)	(2.15)						
Panel B: CH-4+Article_3	factor loading	l (same_article t	vpe)							
	Article_3	Article_6	Article_9	Article_12	Anavst	Industry	Geographic	Technoloav	Statistical	Concept
Article 3		0.931	0.946	0.962	0.977	0.895	0.108	0.654	1.005	0.911
		(40.71)	(28.88)	(22.33)	(14.25)	(13.73)	(4,00)	(8.51)	(9.84)	(02.70)
mktrf		-0.002		-0.017	-0.081	-0.065	0.012	-0.031	-0.074	0000
		-0.002 (_0.23)		(-0.03)		(7 66)	0.012 (0 80)	(UV 1-)	(74.1-)	0.000
			0.001	0.005			(20.0)			0.025
21414		010.0-	100.0-		0.004	-0.014		-0.049		
		(10.0-)	(-0.UZ)	(n. 17)	(/n·n)	(UZU)	(94.0-)	(n/ ·n-)	(77.0)	(0.4U) 0.050
SMB		0.003	0.027	0.04 /	0.048	0.05 (22)	-0.002	-0.029	0.072	0.050
		(0.16)	(06.0)	(1.32)	(0.76)	(0.83)	(/0.0-)	(79.0-)	(0.63)	(/ 9.0)
PMO		-0.001	-0.005	-0.019	-0.064	-0.021	-0.017	-0.049	-0.185	0.093
		(-0.05)	(-0.27)	(-0.75)	(-1.59)	(-0.47)	(-1.05)	(-1.26)	(-2.15)	(1.05)
Adj. R2		0.876	0.823	0.785	0.538	0.492	0.072	0.328	0.361	0.412
# Obs.		464	464	464	464	464	464	464	464	227
Panel A reports the intercept spillover factor. The constructi augmented CH-4 model (i.e., t returns of other momentums c explanatory variables. News cc on the Shenwan-1 classi cation listed stocks on the main boarc is 2012-2020. The sample peri Alphas with t-statistics higher	(or alpha) of r lon of the MS fi he original CH- on MKT , SMB -mention mome 1, while the gec 1, while the gec of of the Shangh od of concept r than 2.00 are h	egressing the tim- actor is given in s 4 model plus a sp , HML, PMO an entum factors are graphic momentum is Au momentum is Au nighlighted in bol	e series of di erei ubsection 4.1. Th ieci c <i>MS</i> factor) d the Article.3 fe based on <i>same.a.</i> am is conducted (e). Shenzhen Stoc d. Panel B report	It long-short mom e column name inco . For example, the rctor. In the last i <i>rticle</i> type under i on the province lev e sample periods i s the factor loads	entum returns of licates the type of a alphas of the ro row of panel A, v identi cation win vel. We do not c frowth Enterpris for other moment of other moment	The CH-4 facto flong-short mom w CH-4+Ariticle we include all the dows including 3 onsider the custo onsider the custo on Market (GEM), tums are 2006-20 um spillover fact	rs (<i>MKT</i> , <i>SMB</i> , <i>HA</i> entum return (I.e., th and indicate the alpha is six momentum spil month, 6-month, 9-r mer momentum due ST shares are exclu 20. Newey and Wes ors on the CH-4+Ar	AL, PMO) (Liu et , the dependent variably is from the time-seri lover factors that at month, and 12-month and 12-month and 12-month ded. The sample per t (1987) adjusted t- iticle_3 model. Coe	al., 2019) plus each e), while the row r es regressions of lo re not related to n h. The industry m ality. The sample riod of news co-me statistics are show cients with t-stat	<i>M S</i> momentum ame indicates the ng-short portfolio ews as additional omentum is based stocks include all stocks include all intion momentum in parentheses.
2.00 al e TIIGIIIIGIIIEA ILI NOIA.										